

THE ROLE OF INTERPERSONAL UNCERTAINTY IN PROSOCIAL BEHAVIOR

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Abstract

In prosocial decisions, decision-makers are inherently uncertain about how their decisions impact others' utility – we call this interpersonal uncertainty. We show that people's response to interpersonal uncertainty shapes well-known patterns of prosocial behavior. First, using standard social allocation decisions, we replicate the classic patterns of ingroup favoritism, merit-based fairness ideals, and self-favoring behavior in dictator games. We then show that these patterns also arise in non-social decisions which have no consequences for others and instead solely reflect responses to interpersonal uncertainty. Behavior is highly correlated across social and non-social decisions, and self-reported interpersonal uncertainty predicts behavior in both situations. Moreover, exogenously varying interpersonal uncertainty shifts prosocial behavior in the direction that avoids such uncertainty. Our results quantify how beliefs in the form of interpersonal uncertainty influence prosocial behavior, which we estimate to be of similar importance to social preferences.

Keywords: prosocial behavior, social preferences, ingroup versus outgroup decisions, dictator games, fairness preferences, interpersonal uncertainty

JEL Classification: C91, D01, D91

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1 Introduction

We as humans can only experience our own utility but not other's utility. Thus, prosocial decisions are inherently decisions made under uncertainty, where we are uncertain about how our decisions impact others' utility. We label this type of uncertainty present in prosocial decisions *interpersonal uncertainty*. If people respond to interpersonal uncertainty as they respond to other types of uncertainties, their response will influence their prosocial decisions. Yet, by and large, theories of prosocial behavior abstract from interpersonal uncertainty, and empirical studies interpret and estimate prosocial behavior assuming certainty.

In this paper, we show theoretically and demonstrate empirically that people's beliefs about and response to interpersonal uncertainty shape behavior across three key paradigms of the social preference literature. Specifically, people's beliefs and their responses to those beliefs reinforce ingroup favoritism in ingroup versus outgroup allocation decisions, self-favoring behavior in dictator games, and redistributive behavior when endowments are earned relative to when received by windfall. Thus, observed prosocial behavior does not solely reflect social preferences. Instead, it reflects a combination of social preferences and beliefs in the form of interpersonal uncertainty, and we demonstrate how to disentangle the two. Our novel explanation highlights the importance of beliefs in driving prosocial behavior, predicts existing evidence on the malleability of prosocial behavior, and informs interventions and policies aimed at changing prosocial behavior (e.g., promoting intergroup contact).

Our preregistered experiments feature a *social* and a *non-social* decision scenario for each of the three aforementioned paradigms of prosocial behavior. The *social* decision replicates a standard decision task used in the literature to elicit the respective preference/behavioral pattern. The *non-social* decision mimics the *social* decision but removes the scope for social preferences while holding the degree of interpersonal uncertainty fixed. Behavior in *non-social* decisions solely reflects responses to interpersonal uncertainty, allowing us to assess its relevance in generating patterns of prosocial behavior.

We illustrate our approach with the ingroup versus outgroup paradigm. In the *social* decision, a decision maker (DM) has to allocate money between two randomly

matched individuals. The two individuals receive the allocated money in the form of gift cards: the decision thus has consequences for both. One of the individuals belongs to the DM's social group, making them an ingroup member, while the other is an outgroup member. Allocating more money to the ingroup member is typically interpreted as an expression of an explicit *preference* or *taste* for the ingroup. For instance, a DM may get a higher marginal utility from the benefit received by the ingroup member (U_{in}) compared to the outgroup member (U_{out}).

We design the *non-social* decisions to rule out any such preference or taste-based channel but retain the interpersonal uncertainty. As before, the DM splits gift card money between an ingroup and an outgroup member, but now without any consequences for either. Instead, the DMs themselves are paid the sum of their matched ingroup and outgroup members' utilities U_{in} and U_{out} from receiving the gift card money. To do so, we measure U_{in} and U_{out} by eliciting the ingroup and outgroup members' willingness-to-pay (WTP) to receive gift card money. The DMs are paid the sum of the two WTPs, weighted by their allocations to the respective members. Importantly, because both WTPs contribute symmetrically to the DMs' total payment, DMs no longer have any taste-based reason to favor either. However, since DMs do not know the WTPs, they face interpersonal uncertainty, and their uncertainty about the distribution of U_{in} and U_{out} might be asymmetric.

In particular, we hypothesize that the DM perceives higher interpersonal uncertainty about U_{out} than U_{in} , for instance, due to her lower familiarity and fewer interactions with the outgroup, and that uncertainty contributes to her ingroup favoritism. As we derive in our theoretical framework, higher uncertainty about the outgroup is indeed sufficient to generate ingroup favoritism under risk aversion. Intuitively, allocations to the ingroup are a "safer bet" and thus preferable under risk-aversion.

We find that behavior in the *non-social* decisions is similar to the *social* decisions. Using shared hobbies/interests, political views, and religious beliefs as groups, subjects allocate on average 61% of the endowment to the ingroup member in the *non-social* decisions compared to 63% in the *social*. Both the average allocations and distributions are similar, as we fail to reject the null of different distributions between *non-social* and *social* for shared hobbies/interests and religious beliefs. Subjects make both *non-social* and *social* decisions in a randomized order. While the last

results were between-subject comparisons obtained using only the first decision, using both for a within-subject comparison reveals a high correlation at the individual level ($r = 0.53$). In fact, the median subject makes the same choice in both decisions. The interpersonal uncertainty inherent in the *social* and *non-social* decisions is thus sufficient to generate the ingroup-favoritism observed in the *social* decisions.

To ensure that the observed similarity between the *social* and *non-social* decisions is not confounded by subjects being confused or inattentive about the incentives, we run three robustness experiments. In these, we systematically vary the *non-social* incentives and find that behavior changes systematically in the directions predicted by our model. For example, ingroup favoritism increases (decreases) significantly when we increase (decrease) the multiplier on the ingroup member's WTP.

Our next treatments directly measure and provide causal evidence for the two-part mechanism suggested in our framework: people perceive higher interpersonal uncertainty about outgroup members and are risk-averse towards this uncertainty. To investigate the first part, we asked subjects to separately state (using Likert scales) how certain they are about ingroup and outgroup members' WTPs. Indeed, subjects perceive significantly higher uncertainty about the outgroup while perceiving no difference in average WTPs. Moreover, higher relative outgroup uncertainty significantly predicts stronger ingroup favoritism in both *social* and *non-social* decisions.

To reveal DMs' attitudes toward interpersonal uncertainty and uncover the causal effect of their attitude on ingroup favoritism, we design a diagnostic treatment where subjects allocate money between individuals randomly chosen from two exogenously created groups. We provide subjects with the actual valuation distribution within each group and exogenously vary this distribution across decisions. Independently, we also vary whether the group members share the DM's social group (ingroup status). The factorial variation of group information and interpersonal uncertainty allows us to isolate the marginal influence of each factor.

In the absence of ingroup/outgroup information, subjects allocate on average 60% of their endowment to the group having a lower variance in valuations, revealing an aversion to interpersonal uncertainty. When we provide both ingroup information and the distribution of valuations, subjects on average allocate 64% to the ingroup when the outgroup's valuations are more uncertain and 49% when the in-

group's valuations are more uncertain. Lastly, when all members across both groups have the same valuation, and thus interpersonal uncertainty is absent, subjects allocate 57% to the ingroup. Accordingly, changes in interpersonal uncertainty shape differences in allocations in the direction predicted by our framework.

When we regress allocations on the treatment variations, we find that the marginal response to interpersonal uncertainty is similar in magnitude to the marginal effect of ingroup preference, with both increasing allocations by around 7% of the endowment. We also estimate a structural model that quantifies the effect of each factor in isolation: aversion towards interpersonal uncertainty is best fit through a CRRA parameter of 0.37 and the strength of pure ingroup preference through a 7% higher allocation to the ingroup. Finally, we use the within-subject structure of our experiment to estimate behavioral types: 33% of subjects respond to interpersonal uncertainty but display no group preferences. 17% of subjects do not respond to uncertainty but display a group preference. In addition, 31% respond to uncertainty and display group preferences, while 20% show neither pattern. Taken together, we find that the majority of subjects respond to interpersonal uncertainty, and their response is quantitatively important in driving observed ingroup favoritism.

We then investigate the importance of interpersonal uncertainty for self versus other behavior using two treatments consisting of a *Self social* and a *Self non-social* decision scenario. While the former is a standard dictator game where DMs allocate gift card money between themselves and another person, the *Self non-social* decision again has no consequence for others. Instead, the DMs' incentives are to maximize the sum of their own WTP and the other person's WTP. Since both WTPs contribute equally to DMs payments, a self-preference no longer predicts more allocation to the self. However, since subjects only know their own WTP but not others' WTP, aversion to interpersonal uncertainty is sufficient to generate "selfish looking choices".

We find that *Self non-social* behavior resembles behavior in *Self social*. Compared to the 69% of the endowment that subjects allocate to themselves in *Self social*, they allocate 64% to themselves in *Self non-social*, and we cannot reject the equality of the two distributions. The similarity extends to within-subject comparisons: the two decisions are highly correlated ($r = 0.71$), making the *Self non-social* decision one of the strongest predictors of dictator game behavior in the literature.

Our first two applications demonstrate how DMs respond when one recipient's (their own or an ingroup member's) valuation-distribution has lower interpersonal uncertainty than the other's. Our third application considers how DMs react when the recipient's valuation-distribution *shifts to the right* (mean-shifted distribution). In particular, we compare dictators who are allocating money from recipient's earned endowment to dictators who are allocating money from a windfall endowment. We hypothesize that dictators believe that the recipient's value-distribution in the first case is a mean-shifted version of the second case, and this contributes to DMs increased hesitancy to take money for themselves in the first case.

We test this hypothesis by designing the *Taking social* decision. In this modified dictator game, instead of the dictator splitting a windfall endowment as in *Self social*, the other person earned the endowment, from which the dictator can take for themselves. This difference has a significant impact on allocations. Subjects in *Taking social* allocate (take) only 41% for themselves, a significant decrease compared to the *Self social* decision. Such behavior has a natural explanation based on merit-based fairness concerns or norms: taking someone's earned money is considered more unfair than keeping money originating from a windfall.

Our *non-social* decision strips the choice of such considerations while retaining its interpersonal uncertainty. Compared to *Self non-social*, in the *Taking non-social* decision, the DM's incentive no longer depends on the recipient's WTP but on the recipient's willingness-to-accept (WTA) to give up gift card money they previously earned, thus capturing the recipient's utility loss from the DM "taking" the gift card money away. Accordingly, DMs split money to maximize the sum of their own WTP and the other individual's WTA. If DMs believe that WTA is higher, something we validate empirically, then the induced *non-social* incentive leads dictators to take less money for themselves. Indeed, subjects allocate 55% to themselves in *Taking non-social*, a significant decrease compared to *Self non-social*, that cannot be explained by fairness concerns. Moreover, allocation choices in *Taking non-social* are significantly associated with taking behavior in *Taking social*. These results suggest that the change in behavior from the dictator game to the taking paradigm is not exclusively driven by fairness considerations but instead is also influenced by the changing utilitarian calculus made under uncertainty.

Related literature. We demonstrate how interpersonal uncertainty influences patterns of prosocial behavior which have been documented across three different strands of the literature. First, a large literature finds evidence that people behave more prosocially towards ingroup members, a finding that is robust across different groups, domains, and methods (such behavior has been labeled ingroup favoritism, parochial altruism or moral universalism, see Charness and Chen, 2020; Shayo, 2020; Enke, 2023, for recent overviews).¹ Second, many studies have documented that in *self versus other* decisions, most people behave prosocially but tend to make choices that favor themselves more than others (see Fehr and Charness, 2023; Capraro et al., 2024, for overviews). Third, in allocation decisions, it has been shown that the source of the endowment matters (see Cappelen et al., 2020, for an overview). In particular, people redistribute less if the money was earned (merited) rather than attained by windfall (Ruffle, 1998; Cherry, 2001; Cherry et al., 2002; Cherry and Shogren, 2008; Oxoby and Spraggon, 2008; Krupka and Weber, 2013). We propose a unified, belief-based mechanism that contributes to each of these patterns of prosocial behavior.² Our results show that due to the inherent presence of uncertainty, observed prosocial behavior cannot be interpreted solely as expressions of social preferences even in standard elicitation tasks. In particular, our results imply that these tasks overestimate the extent of ingroup preferences (or taste-based discrimination), underestimate the degree of altruism in dictator games, and overestimate merit-based fairness preferences. We provide a methodology to separately identify and quantify the roles of beliefs and preferences in driving prosocial behavior.

With our subjective uncertainty-based explanation of prosocial behavior, we relate to a recent literature that explains a range of behavioral patterns through people's cognitive response to (subjective) uncertainty. Enke and Graeber (2023) investigate how people's uncertainty over the optimal decision influences choice under risk, belief formation, and forecasts. In the domain of intertemporal decisions, a series of theoretical studies show that risk and time preferences closely intertwine when DMs

¹See Iyengar et al. (2019) and Böhm et al. (2020) for a review of the recent literature on ingroup favoritism in political science and psychology, respectively.

²Previous belief-based explanations of prosocial behavior have been mainly applied to strategic interactions, such as trust or reciprocity (Berg et al., 1995; Fehr and Gächter, 2000).

are uncertain about future consumption (Sozou, 1998; Dasgupta and Maskin, 2005; Halevy, 2008; Chakraborty et al., 2020) or future preferences (Amador et al., 2006; Chakraborty, 2021). While this literature focuses on characterizing a logical equivalence between subjective uncertainty and risk or time preference patterns, we study the connection between subjective uncertainty and prosocial behavior both theoretically and empirically. In particular, our non-social treatments allow us to quantify the extent to which subjective uncertainty in the form of interpersonal uncertainty drives standard patterns of social behavior.³

To isolate the importance of interpersonal uncertainty in prosocial choices, we construct the non-social decisions by stripping the original social decisions of all other-regarding motivations. This is reminiscent of Oprea (2024) and Enke et al. (2023), who construct diagnostic decisions by stripping risk or discounting-based motivations from standard risky and intertemporal tasks to isolate the role of complexity on decision-making under risk and time.

2 Conceptual framework

Our central premise is that when choosing between different actions of which at least one has consequences for others, a DM perceives *Interpersonal Uncertainty* about how those consequences impact others' utility. That is, she is uncertain about the utility that others receive from the outcome created by her actions.⁴ Specifically, we analyze a DM who has to allocate \$100 between two recipients. In such an allocation decision, interpersonal uncertainty can be understood as the subjective uncertainty the DM perceives about how much utility or valuation each recipient derives from each allocated dollar. We will show that a simple model of interpersonal uncertainty can generate canonical patterns from the literature on prosocial behavior.

³We thus differ from papers investigating prosocial behavior under experimenter-induced *objective risk* over consequences to study ex-post versus ex-ante fairness (e.g., Brock et al., 2013), the use of risk to act selfishly (Exley, 2016), or, risk-induced morality (Chen and Zhong, 2024). Further, Cappelen et al. (2022, 2024) study redistribution decisions when the source of inequality is uncertain. Moreover, Kappes et al. (2018) vary uncertainty about the wealth level of recipients, finding evidence against the hypothesis that people exploit such uncertainty to license more selfishness.

⁴Therefore, it differs from uncertainty about the mapping between actions and outcomes, for instance the uncertainty whether a donation will actually be delivered to a recipient.

2.1 Assumptions about interpersonal uncertainty

For simplicity, we assume that the DM is probabilistically sophisticated and believes that dollars are valued non-negatively. Interpersonal uncertainty then means that the DM believes the per-dollar valuation of recipient j is distributed as $v_j \sim f_j$, where f_j is a probability distribution with non-negative support contained in $[0, b]$, and F_j is the corresponding CDF. For $x \geq 0$, we define $S_j(x) = \int_0^x F_j(y)$.

Next, we assume that these belief distributions have two key features. First, a DM understands that different recipients might derive different values from the same allocated dollar amount based on their personalities, past experiences, socioeconomic status, or tastes. Thus, distributions over the valuations of others are non-degenerate.

Second, the belief distributions for different recipients systematically differ, depending on the DM's familiarity with the recipients, or the source of the \$100 endowment. For instance, suppose one recipient shares their hobbies/ religious/ political interests with the decision-maker (ingroup member) while the other does not (outgroup member). This makes the allocation decision an ingroup versus outgroup tradeoff, which will be our leading example. Facing this tradeoff, a DM might think that shared interests or identity with a recipient is indicative of shared past experiences, economic status, and tastes. As a consequence, DMs may feel less familiarity and thus perceive higher interpersonal uncertainty about the outgroup. Similarly, in situations involving the DM herself as one of the two recipients, DMs may naturally face higher uncertainty about others than about themselves since we are most familiar with our own tastes and circumstances. In other situations, DMs might think that one recipient is systematically more likely to have higher valuations than another recipient. Formally,

Definition 1. *DMs perceive a higher interpersonal uncertainty for recipient 2 than recipient 1 if $S_1(x) \leq S_2(x)$ for all x and $S_1(y) < S_2(y)$ for some y . DMs perceive a mean-shifted interpersonal uncertainty for recipient 2a compared to recipient 2b, if there exists $c \in \mathbb{R}_{++}$ such that for all x , $F_{2a}(x + c) = F_{2b}(x)$.*

The condition for higher interpersonal uncertainty is best understood as a generalization of “ f_2 is a mean-preserving spread of f_1 ” or equivalently “ f_2 is second-order

stochastically dominated by f_2 ”, because the two quoted notions are defined identically with the additional condition that f_1 and f_2 have equal means. We use the concept of higher interpersonal uncertainty to characterize the optimal allocation x^* . In comparison, we use the mean-shift concept to study how x^* changes when the DM’s beliefs about a particular recipient’s valuation-distribution shifts to the right.

2.2 Choice behavior under interpersonal uncertainty

We investigate the case of unbiased utilitarian preferences, which means the utility the DM receives from allocating $x \in [0, 100]$ to the ingroup and $(100 - x)$ to the outgroup is $u_{UTIL} = v_1x + v_2(100 - x)$. As v_1, v_2 are random variables, she maximizes expected utility over the potential utilitarian outcomes:

$$EU(x) = E_{v_1 \sim f_1, v_2 \sim f_2} U(v_1x + v_2(100 - x)) \quad (1)$$

where $U' > 0$ and $E_{v_i \sim f_i}$ denotes the expectation with respect to f_i .

Given this setup, the optimal allocation depends crucially on the response to uncertainty as characterized by U , and on the belief distributions f_1 and f_2 . We will generally assume that $U'' < 0$ which implies that the DM dislikes higher variance over potential utilitarian outcomes. If both f_1 and f_2 are degenerate with different expected values, the DM will allocate 100 to the recipient with the higher expected value.⁵ If both distributions are non-degenerate, Theorem 1 provides the optimal solutions and serves as our prediction for both the *social* and *non-social* decisions we later employ in our experiments.

Theorem 1. *Suppose DM i has unbiased utilitarian preferences and is risk-averse ($U'' < 0$). If f_1 and f_2 are non-degenerate, independent probability distributions, then*

- i) Equal division: If $v_1 \stackrel{d}{=} v_2$ (i.e., $f_1 = f_2$) then i ’s optimal allocation is $x^* = 50$.*
- ii) Ingroup favoritism: If f_2 is a mean preserving spread of f_1 , then i ’s optimal allocation is $x^* \in (50, 100)$.*
- iii) Comparative statics over x^* : Suppose the valuations of the two groups are distributed as v_1 and $c + v_2$ for some constant c and independent random variables $v_1 \sim f_1, v_2 \sim f_2$.*

⁵In the trivial case of degenerate distributions with equal expected value, the optimal allocation is non-unique, as the DM is indifferent between all possible allocations.

Under arbitrary CARA preferences⁶, or under CRRA coefficient < 1 , the optimal allocation satisfies $dx^*/dc \leq 0$.

For the proof, see the Appendix. Part (i) follows from symmetry: a risk-averse DM hedges against interpersonal uncertainty by allocating equally among ex-ante symmetric recipients. (ii) shows that ceteris paribus, if the DM perceives a higher interpersonal uncertainty about one of the recipients, she allocates more to the other recipient. Accordingly, a DM who perceives higher interpersonal uncertainty about the outgroup member will allocate more money to the ingroup, even if they believe that on average, ingroup and outgroup members benefit equally from receiving money. Similarly, a DM who perceives higher interpersonal uncertainty about other's utility than their own utility will allocate more money to themselves, even if they care about others equally and think that on average, everyone benefits equally from money. This motivates our experiments studying the ingroup versus outgroup paradigm and the self versus other paradigm in Sections 3 and 4. Note that in (ii) we use the assumption of equal expected values simply as a benchmark: our key insight is that interpersonal uncertainty can generate ingroup favoritism despite equal expected values.

Finally, part (iii) shows that the DM would decrease the allocation to the ingroup (or the allocation to herself in the dictator game) if her belief about the outgroup's valuation mean-shifts to the right.⁷ For example, if a DM perceives mean-shifted interpersonal uncertainty when allocating a recipient's earned money compared to allocating windfall money (thus, perceiving higher c in the former case), then she would keep less for herself (lower x^*) in the former case. This motivates our experiments studying the giving versus taking paradigm in Section 5.

Will every commonly used welfare criterion deliver the results of Theorem 1 under the right parameters given our assumptions about interpersonal uncertainty? In Appendix C, we show that Rawlsian preferences are insensitive to higher interpersonal uncertainty. We will use this result later in a robustness analysis to show that

⁶For a utility function $U(w)$, the coefficient of absolute risk aversion (ARA) is defined as $r_1(w) = \frac{-U''}{U'}$ and relative risk aversion (RRA) is defined as $r_2(w) = \frac{-wU''}{U'}$. CARA and CRRA imply r_1 and r_2 are constant respectively.

⁷Under extreme risk aversion, when c increases, the marginal return from the states with high v_2 is so low that on the margin, DMs might prefer to allocate more to v_1 to safeguard their utility in the states where v_2 is low.

people respond to our induced incentives in the expected direction.

Relation to preferences. The economic literature generally interprets ingroup favoritism as an expression of ingroup preferences, modeled as a higher utility weight for ingroup compared to outgroup members (e.g., Tabellini, 2008). In psychology, it is often interpreted as an expression of moral values (e.g., Graham et al., 2013). Similarly, various explanations for the fact that people allocate more, but not all of the endowment to themselves in dictator games have been brought forward (e.g., Capraro et al., 2024). Most of these models either implicitly or explicitly assume that DMs weight their own utility differently than others' utility.

However, under interpersonal uncertainty, risk aversion is sufficient to generate ingroup favoritism or self-favoring behavior, differential utility weights are no longer necessary. Importantly, our framework does not imply the absence of social preferences. To the contrary, the described patterns of prosocial behavior emerge precisely because people have social preferences: they care about others' utility, but as this utility is unobserved, they face uncertainty.

3 Ingroup versus outgroup paradigm

We start by studying ingroup versus outgroup decisions before expanding to further prosocial decisions in later sections.

3.1 Experimental design

The experimental sessions using the ingroup-outgroup paradigm (and the other paradigms introduced later) feature two distinct decision situations: *social* decisions, and *non-social* decisions which remove any social preference motivations but retain the inherent interpersonal uncertainty about others' utility.

Ingroup social decisions. For the ingroup versus outgroup paradigm, the *social* decision is a “bystander” money-allocation game – one of the standard experimental decision tasks used to identify differential attitudes towards ingroup and outgroup member (e.g., Chen and Li, 2009; Enke et al., 2022). The game features three individuals, (i) a decision-maker (DM), (ii) one individual who shares a social group

with the DM (ingroup member), and (iii) another individual who is a member of a different group than the DM (outgroup member). The DM is asked to allocate a fixed amount of money between the ingroup and outgroup members. The degree to which DMs allocate more money to the ingroup member reveals their degree of ingroup favoritism.⁸

In total, DMs face three *Ingroup social* decisions, in each allocating \$100 between one ingroup and one outgroup member. Specifically, they allocate money between (i) someone who “shares your interests/hobbies” versus “has different interests/hobbies than you”, (ii) someone who “shares your political views (e.g., a fellow left-winger, or a fellow right-winger, etc.)” versus someone who “has different political views than you” and (iii) someone who “shares your religious beliefs (e.g., a fellow Christian, or a fellow atheist, etc.)” versus someone who “has different religious beliefs than you”.⁹ The allocated money is sent to the ingroup and outgroup member six weeks from the date of the experiment in the form of Amazon gift card money.

Ingroup non-social decisions. Our main contribution is to design and implement a novel decision situation, the *non-social* decision. In this decision, we remove any other-regarding motivations by removing any consequences the DM’s decision has to other individuals. Instead, DM’s choice solely determines their own payoff. DMs split \$100 between an ingroup and outgroup member, using the same groups as in *social*, and the DM’s payoff Π is determined by the following formula:

$$\Pi(x_{in}, x_{out}) = x_{in} \cdot WTP_{in}/100 + x_{out} \cdot WTP_{out}/100.$$

where x_{in} is the money split in favor of the ingroup member, and $x_{out} = 100 - x_{in}$ is the money split in favor of the outgroup member. WTP_{in} and WTP_{out} denote the ingroup and outgroup member’s respective WTP for a \$100 Amazon gift card to be received in six weeks, elicited using a valuation task (explained below). To scale the incentive, the WTP is divided by 100, representing an individual’s WTP per gift card dollar. For example, if the DM split \$40 and \$60 in favor of the ingroup

⁸Particularly, ingroup favoritism is identified independent of the decision-maker’s self-interest. Past research has shown that behavior in such bystander allocation games shows a high test-retest correlation, works equally well when posed hypothetically and incentivized, and is highly correlated with related psychological questionnaires (Enke et al., 2022).

⁹We use the wording of Enke et al. (2022).

and outgroup member respectively, and the elicited WTP of “\$100 Amazon gift card money received in 6 weeks” for the ingroup member were \$80 and for the outgroup member \$60, then the DM’s payoff would be

$$\Pi(40, 60) = 40 \cdot 80/100 + 60 \cdot 60/100 = 68.$$

This payoff function induces utilitarian preferences because we incentivize the DMs to maximize the sum of the WTPs, weighted by the allocations made in their favor. Since DMs do not know the actual WTPs of the matched individuals, such interpersonal uncertainty transforms the *social* decision into an uncertain subjective lottery choice. At the same time, because the WTP is elicited over the same object that is distributed in the *Ingroup social* decision, we keep the degree of interpersonal uncertainty constant between the *Ingroup social* and *Ingroup non-social* decision. Importantly, the ingroup and outgroup member’s WTP enter the utilitarian payoff function symmetrically, so any differences in allocations are driven by differences in uncertainty about the WTPs. We can thus use the comparison of the *Ingroup social* and *Ingroup non-social* decision to assess the relevance of interpersonal uncertainty in driving ingroup favoritism.¹⁰

Valuation task. To elicit the willingness-to-pay (WTP), we use a standard multiple-price-list (MPL). Subjects face a series of binary decisions between (i) receiving a \$100 Amazon gift card in six weeks and (ii) a monetary amount paid today which increased across decisions.¹¹ This procedure reveals the current-day dollar equivalent of receiving gift card money.¹²

Minimizing inattention and confusion. A principal concern when interpreting behavior in the *non-social* decisions is that subjects are inattentive to the incentive structure or misunderstand the parameters of the decision. We employ several measures to mitigate the scope for these confounding factors. First, before completing the

¹⁰The difference between the two decisions can be driven by ingroup preferences, and by the fact that motivated beliefs may come into play in the *social* decision to serve one’s ingroup preferences, but motivated beliefs can only reduce profits made in the *non-social* decisions.

¹¹We enforced single switching by automatically filling the list above and below subjects’ choices.

¹²We used Amazon gift card money because it can be paid anonymously, and the specific gift cards are non-refundable and non-fungible, and because subjects’ valuation differs from the dollar value of the gift card. We implemented the time lag to generate additional variation in subjects’ valuation.

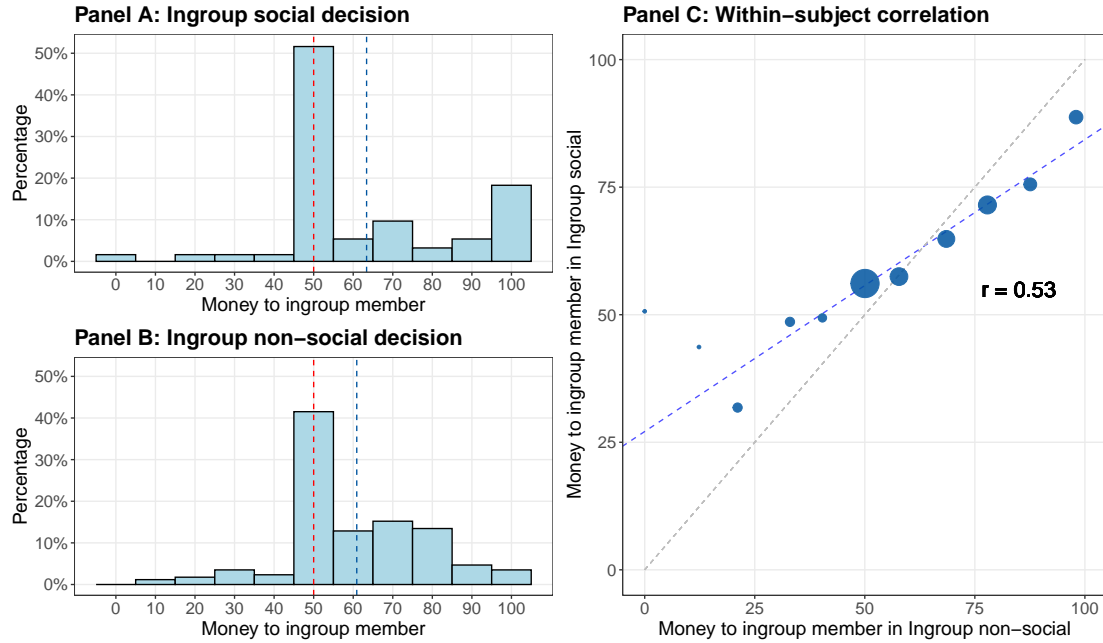
non-social decisions, decision-makers complete the valuation task themselves. That is, they face the WTP elicitation, which familiarizes them with the calculation of the WTP for the incentive. Second, we included several comprehension questions that test whether DM's understood that the *non-social* decisions only have consequences for themselves, not for the other individuals. If they did not answer all questions correctly, we explained them their errors and highlighted the correct answers. This procedure makes it particularly salient that the *non-social* decisions are different from the *social*. Third, to further minimize inattention, we include an explicit disclaimer on the *non-social* decision screens that states "Reminder: your choice only determines your own payment, it does not affect the two individuals." On the decision screen, we also provided DM's with the option to revisit the instructions.

Procedure. We randomized the order of decisions. Half of the decision-makers first face the *social* decision and then the valuation task and *non-social* decision. The other half first face the valuation task and *non-social* decision, and then subsequently the *social* decision. We did not announce beforehand that other decisions would follow the initial decisions, therefore minimizing the scope for contagion from one treatment to the other. This design allows us to analyze within-subject behavior, and compare behavior between-subject by only looking at the first set of decisions.

Data. In total, 119 subjects participated in the ingroup experiment, with 62 subjects first facing the *Ingroup social* decision and 57 subjects first facing the *Ingroup non-social* decision. For this and all further experiments, we used Prolific to recruit online participants living in the US. All experiments were preregistered, see Appendix J for details. We used oTree (Chen et al., 2016) for programming the graphical user interface. Subjects spent a median of 10 to 12 minutes in the experiments and received as compensation the equivalent of an hourly wage between \$10 and \$12 per hour. In each experiment, one randomly selected subject out of the participating subjects had one randomly selected decision implemented with real consequences (between-subject random incentivized system).¹³

¹³In a meta-analysis, Umer (2023) shows that in the context of dictator games paying a subset instead of all subjects does not significantly change behavior.

Figure 1: Main results ingroup versus outgroup decisions



Notes: **Panel A and B:** Histogram of *Ingroup social* (Panel A) and *Ingroup non-social* (Panel B) decisions. The x-axis denotes the amount of gift card money (out of \$100) allocated to the ingroup member instead of the outgroup member. The red line denotes the even split, the blue line the average allocation. In Panel A, the decisions have consequences for the ingroup and outgroup members. In Panel B, the decisions have consequences only for the decision-makers, with their payoff depending on the ingroup and outgroup member’s WTP for the gift card. **Panel C:** Binned scatter plot of *Ingroup social* and *Ingroup non-social* decisions. The blue line displays the linear fit of regressing *Ingroup social* on *Ingroup non-social* decisions. The correlation coefficient is $r = 0.53$. In all panels, the binwidth is 10. Decisions are pooled across the three groups (shared hobbies/interests, political views, and religious beliefs), displaying $n = 186$ (Panel A), $n = 171$ (Panel B), and $n = 357$ (Panel C) decisions by 62, 57 and 119 subjects, respectively.

3.2 Results

We start with the between-subject comparison of the *social* and *non-social* decisions.

Ingroup social decisions. In the *Ingroup social* decisions, subjects allocate on average \$57.48, \$71.05, and \$61.61 to the ingroup member when they share the same interests/hobbies, the same political views, or the same religious beliefs, respectively. In all three cases, we can reject the hypothesis of no ingroup-favoritism ($p < 0.01$, one-sample Wilcoxon tests). Figure 1 panel A displays the distribution pooled over

the three decisions, which replicates the typical distributional pattern found in the literature (e.g., Enke et al., 2022). In 46% of the decisions, subjects display ingroup-favoritism by allocating strictly more than 50% to the ingroup. Outgroup-favoritism is found in 8% of decisions, and in the remaining 46%, subjects allocate 50/50. In total, 73% of subjects display ingroup-favoritism in at least one decision.

Ingroup non-social decisions. Importantly, a similar pattern emerges in the *Ingroup non-social* decisions. Here, subjects allocate on average \$56.86, \$65.02, and \$60.81 when splitting in favor of ingroup members sharing the same interests/hobbies, same political views, and same religious beliefs. As before, we find significant ingroup-favoritism in all three cases ($p < 0.01$, one-sample Wilcoxon tests), even though the decisions have no consequences for ingroup or outgroup members. See Panel B of Figure 1 for the distribution of the pooled decisions. In 61% of the *Ingroup non-social* decisions, subjects display ingroup-favoritism by allocating strictly more than 50% to the ingroup. Outgroup-favoritism is found in 11% of decisions, and in the remaining 28%, subjects allocate 50/50.

Comparing *Ingroup social* and *non-social*. We cannot reject that average ingroup allocations are equal between *Ingroup social* and *non-social* decisions in any of the three cases ($p = 0.59$ for hobbies/interests, $p = 0.22$ for political views, $p = 0.38$ for religious beliefs, unpaired Wilcoxon tests). Further, we cannot reject that the allocation distributions are equal when the groups concern hobbies/interests and religious beliefs ($p = 0.15$ and $p = 0.10$, Kolmogorov-Smirnov test). We can only reject the null of equal distributions in the case of political views ($p = 0.004$, Kolmogorov-Smirnov test). Thus, our *non-social* setup where decisions have no consequences for either group member produces very similar behavior as the standard *social* setup. The most notable difference between the two decisions is the extent to which subjects display maximal ingroup favoritism, i.e., give the entire endowment to the ingroup member. Almost 20% of decisions in *social* display this pattern, mostly stemming from the decisions involving political view groups. In contrast, less than 5% of *non-social* decisions display maximal ingroup favoritism.

Within-subject comparison. Next, we compare behavior between *Ingroup social* and *non-social* on the individual level by including also the second set of decisions of each

subject. We replicate the previously reported between-subjects results also within-subject, see Appendix D.1. Importantly, we find no evidence for order effects, supporting the validity of our within-results (see Appendix E for details). This allows us to correlate behavior in *Ingroup social* with *non-social*. Panel C of Figure 1 displays the distribution of each individual social and non-social decision pair in a binscatterplot. As the figure shows, the two are highly related: ingroup favoritism in *Ingroup non-social* predicts ingroup favoritism in *Ingroup social*, with a correlation coefficient of $r = 0.53$. Therefore, the same subjects that display ingroup favoritism when decisions have consequences for others also display it when their decisions solely affect their own payoff, with the payoff depending on other's WTPs.

Result 1. *We find ingroup-favoritism in Ingroup non-social, which retains interpersonal uncertainty but removes any consequences for ingroup or outgroup members. The distribution of behavior is similar to Ingroup social, which features consequences, and decisions in the two situations are strongly correlated on the individual level.*

3.3 Robustness

Our main results show a high degree of similarity between the *Ingroup social* and *Ingroup non-social* decisions. Next, we present a series of robustness treatments to establish that this similarity is not driven by subjects being confused or inattentive to the experimental design.

3.3.1 Subjects understand and react to *non-social* incentives

If subjects are confused or do not pay attention to the *non-social* incentives, they may treat the *non-social* decisions as *social* decisions. This, in turn, would artificially increase the similarity between the two types of decisions. To test for these confounds, we designed three variations of the *Ingroup non-social* decisions, the *Ingroup incentive*, *Outgroup incentive*, and *Non-social minimum* treatments.

Design. The first two treatments vary the weights put on the ingroup and outgroup members' WTPs in the payoff function used to incentivize subjects' choices. In *Outgroup incentive*, we increase the weight on the outgroup member's WTP to be three

times as high as the ingroup member's WTP:

$$\Pi(x_{in}, x_{out}) = x_{in} \cdot WTP_{in}/100 + 3 \cdot x_{out} \cdot WTP_{out}/100$$

Similarly, in *Ingroup incentive* we increase the weight on the ingroup member's WTP to be three times as high as the outgroup member's WTP:

$$\Pi(x_{in}, x_{out}) = 3 \cdot x_{in} \cdot WTP_{in}/100 + x_{out} \cdot WTP_{out}/100$$

Theoretically, the first incentive induces outgroup favoritism, while the second increases ingroup favoritism relative to the *Ingroup Non-social* incentive, which features symmetric weights. In a third treatment, we instead induce an incentives which according to our framework eliminates favoritism in either direction. Specifically, in the *Non-social minimum* treatment, we incentivize a Rawlsian welfare function. Here, subjects' payoffs are calculated as:

$$\Pi(x_{in}, x_{out}) = \min\{x_{in} \cdot WTP_{in}/100, x_{out} \cdot WTP_{out}/100\}$$

Thus, we incentivize them to choose the allocation that maximizes the utility of the worse-off recipient, irrespective of group affiliation. ¹⁴

Other than the incentives, all aspects of the decisions in the three treatments are identical to the *Ingroup non-social* decisions. Hence, if subjects are inattentive or confused about the *Non-social* incentives so that they erroneously think they face the *Ingroup social* choice instead, we should observe ingroup favoritism in all three treatments.

Results. Compared to an average ingroup giving of \$60.89 across our social groups in *Ingroup non-social* (Section 3.2), subjects give on average \$42.16 to the ingroup member in *Outgroup incentive*, \$67.98 in *Ingroup incentive*, and \$51.31 in *Ingroup non-social minimum*, each time a significant difference (all $p < 0.001$, unpaired Wilcoxon test). Hence, as predicted, ingroup favoritism flips to outgroup favoritism in the first, increases in the second and vanishes in the third case¹⁵. Not only averages,

¹⁴As we show in Appendix C, even if the WTP distribution for the outgroup is a mean-preserving spread of the ingroup's WTP distribution and subjects are risk-averse, the predicted optimal choice under Rawlsian preferences and given our assumptions on interpersonal uncertainty is $x_{in} = x_{out} = 50$, implying no favoritism in either direction.

¹⁵For *Ingroup non-social minimum* we can no longer reject that average ingroup giving is different from the 50/50 split ($p = 0.31$, $p = 0.13$ and $p = 0.95$ respectively for the three social groups,

but also the distribution of choices fundamentally changes, as we show in more detail in Appendix F.1. There, we also analyze the behavior of subjects who are likely inattentive or confused about the incentives, finding that ingroup favoritism is, if anything, less prevalent among those subjects relative to the main experiment.

3.3.2 Results replicate when changing the task determining the incentives

As a final robustness check, we assess the degree to which our results depend on the use of gift card money. Previously, we used gift card money as the good that is allocated in the *social* decisions and as part of the valuation task for the incentives in the *non-social* decisions. In the *Ingroup effort* treatment, we replace the gift cards.

Design. In the *Ingroup effort social* decision, subjects allocate a plain amount of \$10 instead of gift card money between an ingroup and an outgroup member. In *Ingroup effort non-social*, instead of paying subjects based on the sum of the ingroup and outgroup members' WTP for a gift card, subjects are paid based on the members' willingness-to-work (WTW) for a \$10 bonus payment. Eliciting the WTW for the ingroup and outgroup member works as follows: they face a series of binary decisions between (i) completing a number of real effort tasks to receive the \$10 bonus payment and (ii) not working. Each real effort task consisted of moving 30 sliders to the middle position within a 60-second time limit, adopted from Gill and Prowse (2012, 2019). We present the decision in a multiple-price-list format, with the number of completed tasks required for the bonus ranging from 0 to 30 in increments of 2. We define the WTW as the number of tasks for which subjects first switch from preferring to work to not working.

For this robustness exercise, we also change the implementation probability. One might worry that the previous method of implementing the decision of one subject per experiment induces an excessively small winning probability. Therefore, in *Ingroup effort social* and *Ingroup effort non-social*, one out of every ten subjects had one of their decisions implemented. All other elements of the *social* and *non-social* decisions are held constant. In total, 121 subjects participated in this treatment.

one-sample Wilcoxon tests).

Results. We closely replicate our main findings. Again, a majority of subjects displays ingroup favoritism in both decision-situations, and behavior is strongly correlated on the individual level ($r = 0.68$). Our results are not only qualitatively, but also quantitatively very similar, see Appendix F.2 for details.

3.4 Interpersonal uncertainty as mechanism

According to our theoretical framework of Section 2, even if the ingroup and outgroup are perceived as having equal valuations on average, interpersonal uncertainty generates ingroup favoritism through a combination of two factors: first, subjects perceive higher interpersonal uncertainty for the outgroup than for the ingroup members. Second, subjects are averse to higher uncertainty. Accordingly, we first measure the perceived difference in interpersonal uncertainty and correlate it with choice behavior. In a second step, we measure the causal effect of exogenously varied interpersonal uncertainty on choice behavior.

3.4.1 Perceived interpersonal uncertainty differs between groups and predicts behavior

Elicitation. To elicit perceived interpersonal uncertainty, we asked *Ingroup social* DMs the following question after the *social* decisions, separately about their ingroup and outgroup members: “How certain are you about how much the individual (. . .) would value Amazon gift card money?” In the *Ingroup effort social* treatment (Section 3.3.2), “Amazon gift card money” was replaced by “the bonus payment”. Subjects could respond on an 11-point Likert scale from *Very uncertain* to *Very certain*. For the analysis, we re-code the variable so that higher values indicate higher uncertainty. We then create a relative uncertainty measure at the subject level by subtracting every subject’s reported uncertainty over an ingroup member’s valuation from their reported uncertainty over an outgroup member’s valuation.

Results. We find that subjects indeed perceive higher outgroup uncertainty: In *Ingroup social* (*Ingroup effort social*, respectively) the self-reported uncertainty is on average 0.57 (0.95) Likert scale points higher under the ingroup-outgroup classification for different hobbies/interests, 1.01 (1.39) higher for different political views,

and 0.45 (1.01) higher for religious beliefs. All differences are significantly different from zero ($p < 0.05$, paired Wilcoxon tests). Pooled across the three groups, subjects perceive 0.68 (1.12) higher outgroup uncertainty. With respect to the distribution of differences, in 31% (34%) of cases subjects report higher uncertainty for the outgroup; in 12% (5%) they report higher uncertainty for the ingroup and in the remaining 58% (61%) cases, subjects report no difference.

Importantly, differences in uncertainty predict choices: higher uncertainty about the outgroup is associated with stronger ingroup favoritism in *Ingroup social* ($r = 0.30$, $p < 0.001$), *Ingroup effort social* ($r = 0.43$, $p < 0.001$), as well as in *Ingroup non-social* ($r = 0.17$, $p = 0.002$), and *Ingroup effort non-social* ($r = 0.28$, $p < 0.001$). See Figure B.3 in the Appendix for the corresponding binned scatter plots.

Robustness. One might wonder if interpersonal uncertainty, elicited after a *social* or *non-social* choice, could be confounded by the prior choice. For example, subjects might report motivated beliefs that somehow justify their prior choice(s). Hence, in a separate robustness treatment for beliefs (*Ingroup belief measurement*, $n = 59$), we elicited subjects perceived interpersonal uncertainty without them making any *social* and *non-social* choices. We replicate our patterns from the main treatment. Subjects report 0.81 Likert scale points higher uncertainty for individuals having different hobbies/interests, 0.93 for different political views, and 0.66 for religious beliefs (all three significant at $p < 0.01$, paired Wilcoxon tests). Pooled across the three social groups, in 37% of cases subjects report strictly higher outgroup uncertainty, in 58% no difference and in 5% higher uncertainty for the ingroup.

As a validation exercise, at the end of the survey, we also asked subjects to report their perceived interpersonal uncertainty for two artificial groups with exogenously induced objective WTP-distributions, where one WTP-distribution is the mean-preserving spread of the other. Overall, 74% (87% respectively) of subjects report strictly (weakly) higher interpersonal uncertainty about the second group, validating the sensitivity of the Likert-scale measure. See Appendix I for details.

In addition, we asked subjects to report their estimated mean WTP for both ingroup and outgroup members to investigate if mean-differences could also contribute to the observed ingroup favoritism in the *non-social* treatment. On average, subjects

report ingroup members to have a gift card valuation of \$87.85, \$89.64, and \$89.14 for shared hobbies/interests, political views and religious beliefs respectively. For outgroup members, they report valuations of \$88.20, \$86.69, and \$87.85. Thus, on average, subjects' estimates do not differ between ingroup and outgroup ($p = 0.90$, $p = 0.06$, and $p = 0.37$, paired Wilcoxon tests).¹⁶ Pooled across all three groups, subjects believe in 30% of cases that the average ingroup WTP is higher, in 49% of cases they believe them to be equal and in 21% of cases they believe the average outgroup WTP to be higher.

Taken together, these findings demonstrate that subjects indeed perceive higher uncertainty about the outgroup's valuation but do not systematically think that the ingroup has higher valuations on average (no mean shifted distribution).

3.4.2 The causal effect of interpersonal uncertainty on behavior

In our model, risk aversion towards interpersonal uncertainty generates ingroup favoritism whenever the perceived interpersonal uncertainty is higher for the outgroup. In a new treatment (*Ingroup uncertainty*) with 120 subjects, we test this channel by decoupling and independently varying interpersonal uncertainty and ingroup status.

Design. DMs faced seven decisions in random order. In each decision, DMs were endowed with \$10 to allocate between Group A and Group B. Each group consists of two individuals, whom we label as recipients. Recipients were participants of an earlier study in which they provided their willingness-to-work (WTW).¹⁷ DMs were told that after they made their allocation decision, a randomly chosen recipient from each group would receive the money allocated to that respective group. Between decisions, we systematically varied (i) uncertainty over the recipients' WTW within each group and (ii) their ingroup/ outgroup affiliation. Thus, the seven decisions can be classified into four types:

¹⁶Moreover, on average, beliefs are not systematically biased: the average estimates are close to the actual average WTP that is observed in the experiment, which is \$87.

¹⁷Recall that the WTW elicits the willingness to complete tasks to receive a \$10 bonus payment, with tasks ranging from 0 to 30, thus serving as a proxy for the utility value of the bonus payment, measured on an effort scale. To familiarize DMs in *Ingroup uncertainty* with the WTW, they first worked on an example slider task and subsequently faced the WTW elicitation themselves. See Section 3.3.2 for details on the task and elicitation.

1. *Uncertainty without group information* decision: the two recipients of Group A have the same WTW of 12, while the recipients of Group B have a WTW of 4 and 22, respectively. Thus, while the recipient who eventually gets the money from Group A would have a fixed WTW of 12, the recipient receiving the money from Group B could have a WTW 4 or 22. Thus, the DMs face higher interpersonal uncertainty about the WTW of Group B recipients according to Definition 1 in Section 2.
2. *Group information without uncertainty* decision (x2)¹⁸: the two recipients of Group A are ingroup members, while the two recipients of Group B are outgroup members. Moreover, all four recipients have the same WTW of 12.
3. *High uncertainty on ingroup members* decision (x2): the two recipients of Group A are ingroup members, one having a WTW of 4 and the other of 22. The two recipients of Group B are outgroup members, both having a WTW of 12. Hence, the ingroup has a higher WTW variation.
4. *High uncertainty on outgroup members* decision (x2): the two recipients of Group A are ingroup members, both having a WTW of 12. The two recipients of Group B are outgroup members, one having a WTW of 12 and the other of 22. Hence, the outgroup has the higher WTW variation.

The *uncertainty without group information* treatment reveals DMs' attitude towards higher uncertainty about WTWs without the confound of ingroup preferences. Because the expected value of WTWs is higher in Group B¹⁹, DMs who still allocate more to Group A reveal their aversion to the uncertainty in WTWs, and hence their risk aversion. The *group information without uncertainty* decision, on the other hand, reveals ingroup favoritism in the absence of interpersonal uncertainty. The last two decision-situations reveal the extent to which interpersonal uncertainty influences ingroup favoritism. When presenting the WTW and social group information, we

¹⁸The "x2" indicates that there were two decisions of this type, one involving shared hobbies and the other involving shared political views as ingroup. We did not include religious beliefs because, based on previous results, we expected similar behavior and wanted to avoid adding more decisions.

¹⁹Intrinsically, this argument holds when recipients' value of \$10 is directly equal to their WTWs, as well as when recipients' valuation is equal to an increasing and convex disutility function of the WTW. The latter assumption is well supported by empirical evidence on real effort tasks (see e.g., Gill and Prowse, 2019).

randomized the group's position on the screen (left or right) and which information was presented first, balancing the presentation of the two pieces of information.

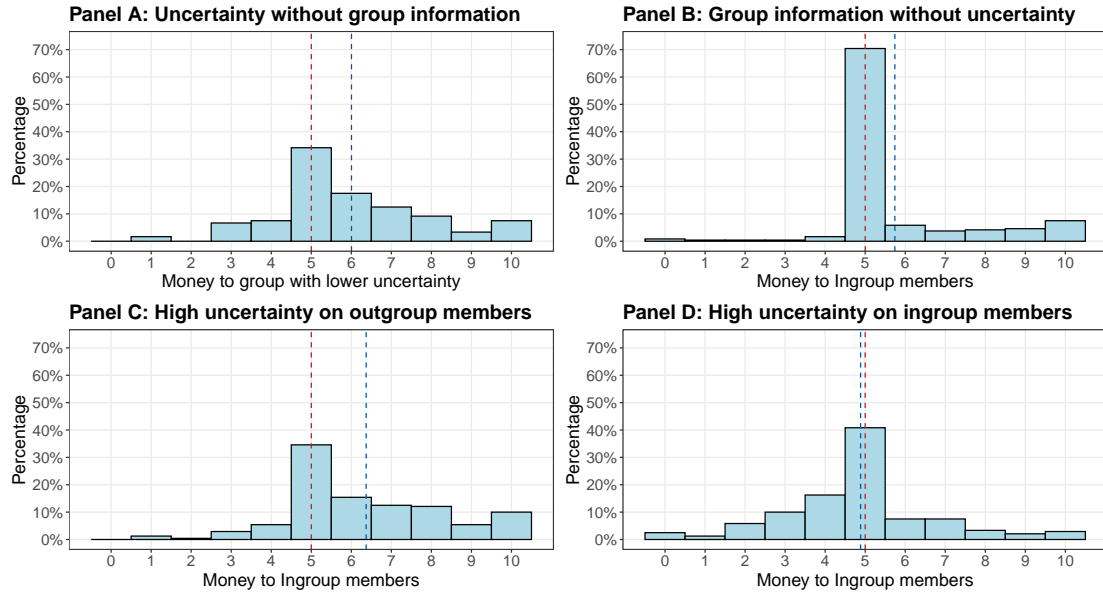
Importantly, the WTW information is based on recipients' *stated* WTW, not their actual exerted effort. While facing the WTW elicitation, recipients knew that with a 50% chance they would complete their selected WTW tasks and with a 50% chance they would not work but later potentially receive money from other participants (the DMs) who observe their WTW. We then only matched the latter half of subjects with the DMs of the *Ingroup uncertainty* experiment. Thus, none of the recipients had worked based on their WTW choices, a fact we saliently communicated to the DMs. This ensures that the DMs have no fairness reason to be partial to one recipient over the other based on exerted effort.

Results. Figure 2 displays the distribution of choices. Starting with Panel A, we observe that the majority of subjects are risk-averse towards interpersonal uncertainty even in the absence of group information: 54% allocate more than 50/50 to the group with lower WTW variance, 27% allocate 50/50 and a minority of 19% allocate more to the group with the higher WTW variance. On average, subjects allocate \$6.00 to the group with lower WTW variance, which is significantly more than the 50/50 benchmark ($p < 0.001$, one-sample Wilcoxon tests).

In the case of all recipients having the same WTW (Panel B), subjects allocate on average \$5.74 to the ingroup (pooling the decision across both social groups), which is significantly different from the 50/50 split ($p < 0.01$, one-sample Wilcoxon test). This quantifies the extent to which ingroup preferences drive ingroup favoritism since interpersonal uncertainty is absent in this decision.

Panels C to D then document how interpersonal uncertainty being higher in the ingroup or the outgroup changes the magnitude of ingroup favoritism. The amount allocated to the ingroup increases from \$5.74 to \$6.37 when the outgroup is more uncertain (has higher variation in WTWs), and decreases to \$4.89 when the ingroup is more uncertain (both $p < 0.001$ compared to the no uncertainty decision, paired Wilcoxon tests). Moreover, changing the uncertainty changes the entire distribution of choices. When the outgroup has the higher WTW variance, the modal (62%) DM displayed ingroup favoritism, while 12% displayed outgroup favoritism and 26% had

Figure 2: The causal effect of interpersonal uncertainty on allocation decisions



Notes: **Panel A:** The x-axis denotes the amount of money (out of \$10) allocated to the group with a lower variation in willingness-to-work (WTW). In **Panels B to D**, the x-axis denotes the amount allocated to the group whose members share a social group with the decision-maker instead of the group whose members are from a different social group. In Panel B, all group members have the same WTW. In Panel C, the outgroup has the higher WTW variation. In Panel D the ingroup has the higher variation. The red line denotes the even split, the blue line the average allocation. In all panels, the binwidth is 10. Decisions are pooled across two social groups (shared hobbies/interests and political views), displaying $n = 120$ (Panel A) and $n = 240$ (Panel B to D) decisions made by 120 subjects.

the 50/50 split. Once the recipients of both groups have equal WTWs, the model subject (66%) now chooses the 50/50 split. Here, only a minority of 29% and 5% strictly favor the ingroup and outgroup, respectively. Lastly, when the ingroup has the higher WTW variance, the modal subject (40%) switches to outgroup favoritism, while 31% of choices show ingroup favoritism and 29% are 50/50 choices. These results establish that responses to interpersonal uncertainty causally influence the extent of ingroup favoritism.

Information choice. We further included an information decision after subjects completed the seven allocation decisions. In this eighth decision, subjects did not have information about group membership and WTW, and could choose to learn about one of the two before allocating the money. The decision thus reveals which of the

two is of primary importance to the subjects' decision process. In total, 81% of subjects chose to learn about WTW variation, indicating a preference for information about uncertainty over social group membership. For details, see Appendix G.

3.5 The quantitative importance of interpersonal uncertainty

Next, using the data from the previous section, we quantify the relative importance of interpersonal uncertainty and ingroup preferences in driving ingroup favoritism, using a reduced-form analysis and a structural model of prosocial decision-making.

3.5.1 Reduced form analysis

Average ingroup favoritism. To estimate the effects of uncertainty and preferences on ingroup preferences, we use the following model: $alloc_{i,d} = c_0 + c_1unc_d + \varepsilon_{i,d}$, in which $alloc_{i,d}$ denotes the allocation to the ingroup by individual i in decision d . We normalize the variable by subtracting 5 from the actual giving so that the 50/50 split benchmark implies $alloc_{id} = 0$. The variable unc_d is equal to 1 for decisions with *Higher WTW uncertainty on outgroup members*, equal to -1 for decisions with *Higher WTW uncertainty on ingroup members*, and equal to 0 when uncertainty is absent for both groups. Thus, c_1 measures how much the allocation is affected by uncertainty in either direction. Accordingly, ingroup preferences are measured by the constant c_0 that captures how much more subjects allocate to the ingroup on average in the absence of interpersonal uncertainty ($unc_d = 0$).

We find that both ingroup preferences and interpersonal uncertainty significantly influence on behavior, and their influence is *nearly equal in magnitude*. While the coefficient on preferences is 0.741, the coefficient on uncertainty is 0.742, both being significant at the 0.001 level (see Appendix Table A.1 column 1). We also estimate an alternative specification where we split the *Interpersonal uncertainty* variable into two indicators for the two uncertainty decisions. Here, we find that the influence of adding uncertainty on the ingroup is slightly larger than adding uncertainty on the outgroup (see Appendix Table A.1 column 2).

Subject-level type analysis. We can further exploit the within-subject structure to identify distinct behavioral types. We say a subject reveals a *group-based preference*

if they choose a different allocation than 50/50 in *at least one* of the two decisions that provided *group information without uncertainty*. We say that a subject *responds to interpersonal uncertainty* if, for both social groups (interest and political views), their *outgroup uncertainty* decision was different from their *ingroup uncertainty* decision.²⁰ With this categorization, we find that 20% of subjects neither respond to uncertainty nor display a group preference. 33% of subjects respond to uncertainty but do not display a group preference, while 17% do not respond to uncertainty but display a group preference. Finally, 31% both respond to uncertainty and display a preference. Hence, interpersonal uncertainty is relevant for 64% of all subjects, while group preferences are relevant for 48%.

3.5.2 Structural model

Setup. Suppose the representative DM has to distribute M units between two individuals 1 and 2, whose group identity can take one of three values: $G(i) \in \{in, out, \emptyset\}$, where \emptyset means unknown. Suppose the DM believes that the valuations of money received by the two individuals 1 and 2 are distributed as f_1 and f_2 respectively, and suppose $x_{IU}(f_1, f_2, \gamma)$ is the choice that maximizes the expected (utilitarian) utility (see equation 1) given the two distributions. To parameterize risk aversion, we assume CRRA utility $u(w) = \frac{w^{1-\gamma}}{1-\gamma}$, with γ as the risk aversion parameter. In our experiment, under the assumption that the valuation of money is measured as WTW, f_1 and f_2 is either the degenerate lottery $L_1 = (12, 1)$ (tasks) or the 50-50 lottery $L_2 = (h = 22, 0.5; l = 4, 0.5)$. Under these distributions and under CRRA risk-preferences, x_{IU} has the following closed form solution:

$$x_{IU}(f_1 = (m, 1), f_2 = (h, 0.5; l, 0.5), \gamma) = \frac{M \cdot h \left(\frac{m-l}{h-m}\right)^{\frac{1}{\gamma}} - M \cdot l}{(m-l) + (h-m) \left(\frac{m-l}{h-m}\right)^{\frac{1}{\gamma}}}$$

²⁰Thus, we use a more conservative identification criterion for the response to uncertainty, because we require subjects to respond to uncertainty across both pairs of choices. In contrast, for identifying group-based preferences only one choice needs to be different from the “no favoritism” benchmark. In total, 79% of subjects respond to at least one change in uncertainty from ingroup uncertainty to outgroup uncertainty.

Finally, we assume that the DM's optimal allocation to individual 1 in observation j is as follows:

$$x_{1j} = \begin{cases} \frac{M}{2} + b + \varepsilon_j & \text{if } f_1 = f_2, G(1) = in, G(2) = out \\ x_{IU}(f_1, f_2, \gamma) + \varepsilon_j & \text{if } f_1 \neq f_2, G(1) = G(2) \\ a_{IU} \cdot x_{IU}(f_1, f_2, \gamma) + a_{ING} \cdot \left(\frac{M}{2} + b\right) + \varepsilon_j & \text{if } f_i \neq f_o, G(1) = in, G(2) = out \end{cases}$$

In the first case of symmetric interpersonal uncertainty, the ingroup preference factor b alone determines the allocation. The normal noise parameter $\varepsilon_j \sim N(0, \sigma^2)$ is i.i.d across observations. In the second case of symmetric group information, interpersonal uncertainty alone determines the final allocation. When we have a conjunction of the former two cases, the optimal allocation combines the influence of both factors: $a \lesseqgtr 1$ quantifies if the influence of the corresponding channel shrinks ($a < 1$), stays unchanged ($a = 1$), or expands ($a > 1$) when both factors are present.

Results. We jointly estimate $\gamma, b, a_{IU}, a_{ING}, \sigma$ to maximize the likelihood of the observed data. We estimate a CRRA parameter of $\gamma = 0.374^{21}$ and the extent of pure ingroup preference to be $b = 0.741$. We estimate a weight of $a_{IU} = 0.739$ on interpersonal uncertainty and a weight of $a_{ING} = 0.336$ on ingroup preferences. Thus, when both factors operate simultaneously, the estimated influences of interpersonal uncertainty and ingroup preferences diminish to 74% and 34% of their respective influences when they operated in isolation. For more details on the estimation, see Appendix Table A.2.

This sub-additive feature ($a_{IU}, a_{ING} < 1$) helps one interpret our results from the *social* and *non-social* decisions in light of the reduced-form results from the *Ingroup uncertainty* treatment. One might (incorrectly) think that the quantitative similarity between the *social* and *non-social* decisions implies that interpersonal uncertainty is sufficient to explain *all* observed prosocial behavior, which would be at odds not only with the previous literature but also with our results in the *Group information without uncertainty* decision. However, the sub-additivity feature explains that when

²¹This implies a higher risk-aversion towards interpersonal uncertainty than towards monetary risk. For comparison, across 16 studies employing the Gneezy-Potters (Gneezy and Potters, 1997) investment task over money, which our setup mimics, Crosetto and Filippin (2016) report an average CRRA parameter $\gamma = 0.3$.

one compares a treatment where both factors are present to one where only interpersonal uncertainty matters, the influence of interpersonal uncertainty in the latter treatment expands and thus partly compensates for the lack of group preferences.

4 Self versus other paradigm (Dictator game)

Our experimental design naturally extends to choices involving tradeoffs between one's own utility versus the utility of others (self versus other decisions), as does the idea that interpersonal uncertainty shapes behavior in these tradeoffs.

4.1 Design

Similar to the ingroup versus outgroup case, DMs face a *Self social* and a *Self non-social* decision, in randomized order. Before the *Self non-social* decision, they also complete the valuation task for \$100 Amazon gift card money received 6 weeks later.

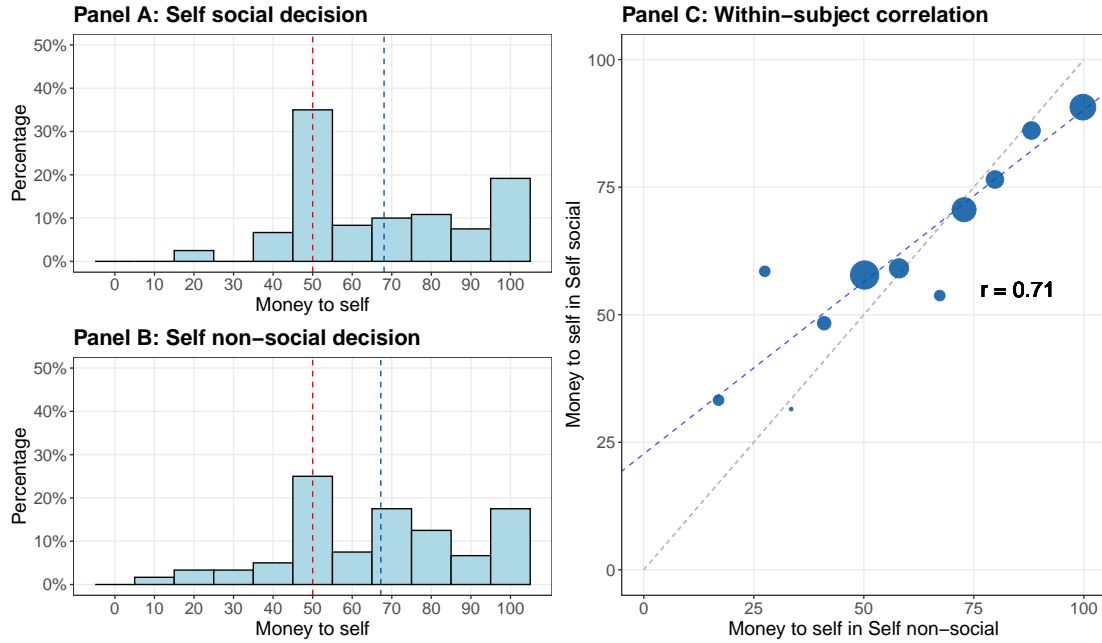
Self social decision. For the *Self social* decision, we endow decision-makers with \$100 which they can allocate between themselves and another individual they have been matched with (without any information about group affiliations). The allocated money is paid out to both parties in the form of Amazon gift card money, six weeks from the date of the experiment. Hence, the *Self social* decision is the standard dictator game: it has consequences for the DM as well as the other individual.

Self non-social decision. In the *Self non-social* decisions, decision-makers split \$100 between themselves and another individual, and we remove any social consequences like we did in *Ingroup non-social* decisions. That is, neither the DM, nor the matched participant receive the money that is split. Instead, only DM's themselves receive a reward based on the following formula:

$$\Pi(x_{self}, x_{other}) = x_{self} \cdot WTP_{self}/100 + x_{other} \cdot WTP_{other}/100$$

where x_{self} and x_{other} are the amounts allocated to self and to the matched individuals respectively, and WTP_{self} and WTP_{other} are their respective WTP for the gift card. Decision-makers are thus incentivized to maximize the sum of their WTP and the WTP of the other individual they are matched with, with both WTPs receiving

Figure 3: Main results self versus other decision



Notes: **Panel A and B:** Histogram of the *Self social* (Panel A) and *Self non-social* (Panel B) decision. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to themselves instead of another individual. The red line is the even split, the blue line the average allocation. In Panel A, the decision has consequences for the subjects and the other individuals. In Panel B, the decision has consequences only for the subjects, with their payoff depending on their and the other individual's WTP for the gift card. **Panel C:** Binned scatter plot of the *Self social* and *Self non-social* decision. The blue line displays the linear fit of regressing the *Self social* on the *Self non-social* decision. The correlation coefficient is $r = 0.71$. For all three panels, the binwidth is 10. Displayed are $n = 61$ (Panel A), $n = 59$ (Panel B) and $n = 120$ (Panel C) decisions.

equal weight. All other elements match the ingroup versus outgroup setting. In total, 120 subjects faced the *Self social* and *Self non-social* decisions.

4.2 Results

***Self social* decision.** In *Self social*, subjects allocate on average \$69.05 to themselves, thereby allocating significantly more to themselves compared to the equal split ($p < 0.001$, one-sample Wilcoxon test). Figure 3 panel A displays the distribution, which replicates the typical pattern of dictator game behavior found in the literature (Engel, 2011). In total, 61% of subjects allocate more money to themselves, 3% allocate more

to the other person, and 36% implement the 50/50 split.

Self non-social decision. In *Self non-social*, subjects allocate on average \$64.12 to themselves, again a significant deviation from the equal split ($p < 0.001$, one-sample Wilcoxon test). As Figure 3 panel B shows, the distribution is also similarly shaped as in the *Self social* case. In total, 58% of subjects allocate more money to themselves, 17% allocate more to the other person, and 25% implement the 50/50 split.

Comparing *Self social* and *non-social*. Allocations in the *Self non-social* setting closely replicate the behavior we observe in *Self social*. Statistically, we cannot reject that the average amount that subjects allocate to themselves is equal across *social* and *non-social* decisions ($p = 0.27$, unpaired Wilcoxon test). Similarly, we cannot reject that the distribution of allocations is equal across the two decisions ($p = 0.39$, Kolmogorov-Smirnov test). These results also replicate within-subject. Figure 3 panel C binscatter-plots each individual's *social* and *non-social* decision pair. The two decisions are highly correlated at the individual level, with a correlation coefficient of $r = 0.71$. Hence, the *Self non-social* decision strongly predicts the *Self social* decision.

Robustness. To show that people are attentive to the *non-social* incentives in the self versus other setting, we use the same the incentive treatment as in the ingroup versus outgroup setting (Section 3.3) and apply it to the current setting. Our results are similar: subjects understand the incentives and react to them as hypothesized. We provide the results in Appendix Section H.

Result 2. *Self non-social choices replicate the self-favoring behavior found in Self social choices. The distributions are similar and strongly correlated on the individual level.*

4.3 Relating interpersonal uncertainty to dictator game giving

After the *Self social* decision, subjects reported self-reported interpersonal uncertainty on a 0 to 10-point Likert-scale both over their own valuation and their perception of the other individual's valuation. Subjects report on average 2.59 Likert-scale points higher uncertainty for the other person's valuation than their own, which is again significantly different from the no difference benchmark ($p < 0.001$, paired Wilcoxon tests). In total, 72% of subjects report a higher uncertainty about the

other person's valuations, 6% report a higher uncertainty about their own valuation, and 23% report equal ratings. Importantly, the difference in ratings again predicts choice behavior. Higher relative uncertainty about the other individual is associated with subjects allocating more money to themselves, e.g., displaying less other-regarding behavior in the *Self social* ($r = 0.35, p < 0.001$) and *Self non-social* decision ($r = 0.24, p < 0.01$). See Appendix Figure B.4 for a binned scatter plot.

Robustness. To ensure that our uncertainty measure was not confounded by prior decisions²², we conducted a robustness treatment (*Self/other belief measurement*, $n = 61$) where subjects only reported their uncertainty about own and other's valuation on the 11-point Likert-scale. We replicate the previously reported results: subjects perceived 3.11 Likert-scale points higher uncertainty over the other person's valuation relative to their own ($p < 0.001$, paired Wilcoxon test). In total, 84% of subjects report higher uncertainty about others' valuations, 10% report the same degree of uncertainty, and 7% report higher uncertainty about their own valuation.

We also asked them to estimate the mean WTP of other subjects. On average, subjects report others to have a mean WTP of \$83.43 and themselves have a WTP of \$86.69, on average. Thus, on average, subjects believe others to have a lower WTP than themselves ($p = 0.03$, paired Wilcoxon test). Overall, for 52% of subjects, their own WTP is higher than their estimate of others' average, the reverse is true for 41%, and for 7% they coincide. Accordingly, if the modal subject was a pure utilitarian who altruistically treats others equally as himself, he would still display self-favoritism. This result echoes our central argument that not accounting for interpersonal uncertainty in self-favoritism might result in underestimating the extent of altruism.

5 Giving versus taking paradigm

Next, we turn to studying how mean-shifted interpersonal uncertainty influences redistribution behavior. The previous literature primarily finds that redistribution behavior is merit-based: people redistribute less from initial endowments if these en-

²²For instance, motivated beliefs that excuse more giving to the self in the *social* decision, can also influence the uncertainty measure thereafter.

dowments are earned compared to generated by chance (Cappelen et al., 2020). In particular, in the context of dictator games, several studies (discussed in the introduction) show that if the initial endowment was earned instead of being windfall, then dictators increase their allocation towards the individual earning the endowment.

This behavior is typically attributed to fairness preferences (e.g., Tungodden and Cappelen, 2019), fairness-based social norms (Krupka and Weber, 2013), or the role of property rights (Oxoby and Spraggon, 2008). We offer an alternative explanation: if people perceive that on average, the disutility from losing earned money exceeds the utility from gaining money (i.e., a gain-loss asymmetry), then dictators would perceive mean-shifted interpersonal uncertainty for recipients who have earned the endowment compared to recipients who have not. Then, based on our framework and our result (iii) of Theorem 1, a simple utilitarian motive under uncertainty would also lead to the same asymmetry between giving and taking environments. Our next treatments test this channel of mean-shifted interpersonal uncertainty.

5.1 Design

Following the typical setup of the literature, we alter our dictator game from a giving environment to a taking environment. DMs face a *Taking social* decision and a *Taking non-social* decision. In both *Taking* decisions, DMs are matched to a previous participant who has earned \$100 for participating in a previous study, scheduled to be paid in 6 weeks from the study day. In total, 123 subjects participated in this experiment.

Taking social decision. In the *social* variant, the DM decides whether to take some or all of the money that the other participant has earned for themselves, adapting the design of Oxoby and Spraggon (2008). The chosen allocation is then implemented with consequences for the DM and the other participant.

Taking non-social decision. In the *non-social* variant, we replicate the setup described in Section 3 with one key difference: because the other participant already earned the \$100 that was up for splitting, the DM's utilitarian incentives were calculated using the other participant's willingness-to-accept (WTA) for gift card money, instead of their WTP. Thus the DM's payment depended on their own WTP and the

matched participant’s WTA. Specifically, the incentive for the DM is as follows:

$$\Pi(x_{self}, x_{other}) = x_{self} \cdot WTP_{self}/100 + x_{other} \cdot WTA_{other}/100$$

with x_{self} and x_{other} denoting the money DMs allocate to themselves and the other individual respectively, WTP_{self} is their own WTP and WTA_{other} is the other individual’s WTA for the gift card money.

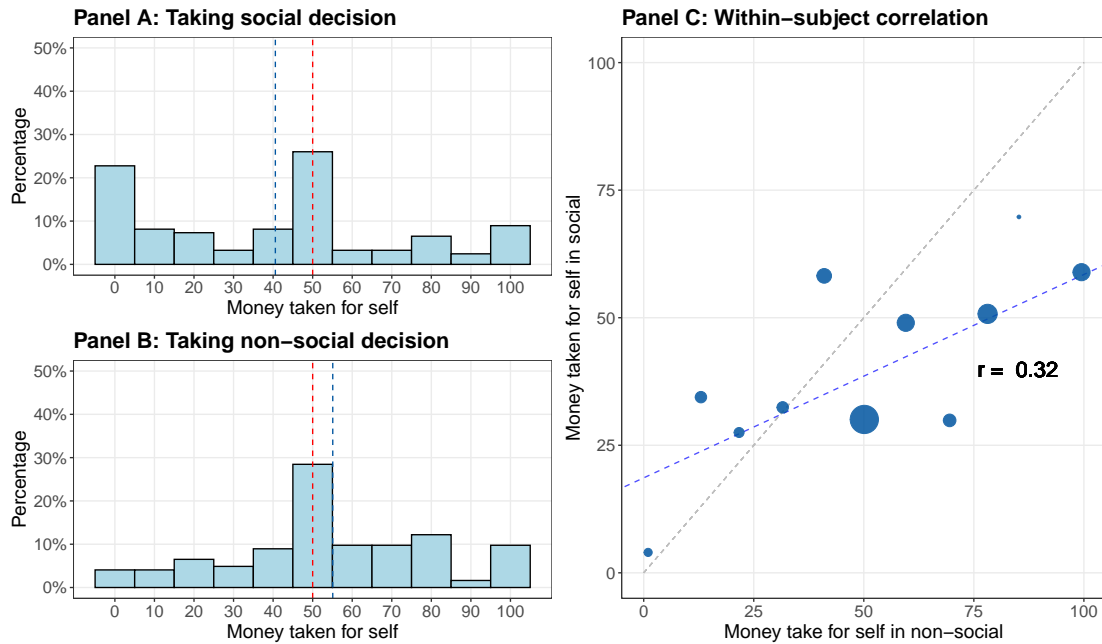
After the DMs participated in the MPL that elicits their WTP, we explained to them the following details about matched participants: First, the matched participants earned the \$100 gift card that would pay in 6 weeks, through their participation. Then, we asked them whether they would be willing to give away the gift card in exchange for an immediately payable monetary amount. We ask this question for different amounts of the immediately payable money, using an MPL, to elicit their WTA. The DMs are already familiar with the MPL-elicitation method at this point. We emphasize to DMs that the only difference between their’s and the matched participant’s elicitation is, instead of having the option to receive the gift card, the matched participants already ‘owned’ the gift card and had the opportunity to sell it.

Multiple studies have found that WTA is, on average, higher than WTP (e.g., Camerer, 1995), and hence $WTA > WTP$ is a well-established empirical pattern. Our central hypothesis is that, if DMs also anticipate the WTA-WTP gap as mean-shifted interpersonal uncertainty, then utilitarianism provides a novel foundation for differences in giving and taking (Theorem 1). In particular, under $WTA > WTP$, we predict that the amount allocated to the matched participant should increase (compared to the giving paradigm) not only in *Taking social*, but also in *Taking non-social*. Further, because *Taking non-social* does not feature any scope for fairness attitudes, we can separate our channel from a fairness channel.

5.2 Results

Comparing *Taking* to the *Giving* setting. We find that subjects allocate significantly less money to themselves in the *Taking* compared to the *Giving* setting of Section 4. Comparing *Self social* with *Taking social* decisions, we see a significant decrease of \$27.48 in the amount subjects allocate to themselves, using the within-subject data controlling for the order (see Appendix Table A.3 column (1) for details). We thus

Figure 4: Giving versus taking results



Notes: **Panel A and B:** Histogram of the *Taking social* (Panel A) and *Taking non-social* (Panel B) decision. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to themselves instead of another individual. The red line denotes the even split, the blue line the average allocation. In Panel A, subjects' decision has consequences for themselves and the other individual. In Panel B, the decision has consequences only for the subjects, with their payoff depending on their WTP and the other individual's WTA for the gift card. **Panel C:** Binned scatter plot of the *Taking social* and *Taking non-social* decision. The blue line displays the linear fit of regressing *Taking social* on the *Taking non-social* decision. The correlation coefficient is $r = 0.32$. For all three panels, the binwidth is 10. Displayed are $n = 123$ decision-pairs by 123 subjects.

replicate the common finding of aversion to taking from earned endowments in the literature with our *social* decisions. When comparing *Self non-social* with *Taking non-social* decisions, we also find a significant decrease of \$12.08 in the amount subjects allocate to themselves (Appendix Table A.3 column (2)). Therefore, incentivizing DMs with the other individuals' WTA instead of their WTP induces DMs to allocate less to themselves. Yet, the decrease in allocation to the self from the *Giving* to the *Taking* setting is smaller in the *non-social* case compared to the *social* case.

Comparing *Taking social* and *Taking non-social*. Figure 4 displays the comparison. Panel A shows the distribution of choices in the *Taking social* decision, where 26% of

subjects allocate more money to themselves, 51% allocate more to the other person, with the remaining 23% allocating the even split. In the *Taking non-social* decision, displayed in Panel B, 46% allocate more money to themselves, 29% allocate more to the other person, and 24% split evenly. We see a significant within-subject correlation of $r = 0.32$ ($p < 0.001$) between *Taking social* and *Taking non-social* (Panel C of Figure 4). Thus, taking behavior in the *social* decision correlates with the *non-social* decision that does not feature taking (not even in how the instructions were framed). Contrary to the other settings, these decisions differ in average allocations ($p < 0.001$, unpaired Wilcoxon test), and distributions ($p < 0.01$, Kolmogorov-Smirnov test). These results suggest that *social* decisions are also driven by motives that are absent in the *non-social* decisions.

A potential motive comes from the observation that 22% of subjects choose to take \$0 for themselves in *Taking social*, while only 3% do so in *Taking non-social*. In contrast, in *Self social* and *Self non-social*, not a single subject chooses to give everything to the other individual. This pattern suggests that fairness preferences are also at work, e.g., some subjects have a strong libertarian fairness view (Alma et al., 2020) or adhere to a deontological motive that it is not permissible to take money from someone (Bénabou et al., 2024). Interestingly, those subjects refusing to take any money completely explain the gap between *Taking social* and *Taking non-social*. If we focus only on subjects who take more than \$0 for themselves in *Taking social*, we can no longer reject the equality of average giving between *Taking social* and *Taking non-social* ($p = 0.22$, unpaired Wilcoxon test)²³ and distributions become more similar ($p = 0.09$, Kolmogorov-Smirnov test), see Appendix Figure B.5. Similarly, the within-subject correlation increases to $r = 0.45$.

A necessary condition for interpreting our hypotheses and results is that subjects indeed perceive a positive difference in the utility impact of taking earned money and giving windfall money, i.e., in other's WTA and WTP. To validate this assumption, we asked subjects whether they generally think that a person's WTA for the gift card is higher, lower or equal to the WTP. In total, 46% of subjects believe WTA to be higher than WTP, 29% believe WTP to be higher, and 24% believe both to be equal. Thus,

²³Note that this effect is not mechanical because we remove both *social* and *non-social* decisions due to the within-subject structure of our data.

subjects believe $WTA > WTP$ on average.

Robustness. For robustness, subjects in the *Self/other belief measurement* treatment ($n = 61$), who reported the mean and the uncertainty over other's WTP valuations, also reported the same quantities about other's WTA valuations. Thus, each subject reported their estimated mean and their perceived uncertainty over other's WTP/WTA. On average, subjects reported a \$8.00 higher mean WTA than mean WTP (\$91.43 compared to \$83.43), a significant difference ($p < 0.001$, paired Wilcoxon-test). In total, 64% of subjects estimated the mean WTA to be higher than the mean WTP, 20% the estimated the reverse, and 16% estimated both to be equal.

With respect to uncertainty, subjects report a 0.51 Likert scale points higher uncertainty about others' WTP than about others' WTA, a significant difference ($p = 0.02$, paired Wilcoxon-test). These results suggests that both mean-shifted interpersonal uncertainty and differences in perceived interpersonal uncertainty contribute to the giving versus taking differences.

Result 3. *Subjects allocate more money to the other person when allocating the other person's earned money (Taking social) than when allocating windfall money (Self social). The allocations are ranked similarly in Self non-social with Taking non-social.*

6 Conclusion

In this paper, we provide a conceptual framework and implement a series of experiments documenting how interpersonal uncertainty bolsters ingroup favoritism, weakens altruistic giving, and shapes redistributive behavior. We show that a significant degree of heterogeneity in prosocial behavior, both within a given decision setting and between different settings, is driven by people's differential response to interpersonal uncertainty. As a consequence, precise identification of social preferences from prosocial behavior requires explicit accounting for interpersonal uncertainty. Otherwise, depending on the nature of interpersonal uncertainty, parameters of social preferences may be over- or underestimated. We also demonstrate an experimental design to disentangle uncertainty from preferences: a researcher can exogenously vary interpersonal uncertainty to explicitly measure and control for it. For

instance, in our experiment, we provide subjects with information so that interpersonal uncertainty switches between recipients, or is balanced among recipients.

Finally, our framework supports the idea that prosocial behavior is malleable. Under the assumption of "exposure reduces interpersonal uncertainty", it helps explain the dynamics of prosocial behavior in response to intergroup contact created by spatial proximity (Bursztyn et al., 2024), shared classes (Rao, 2019), shared living (Corno et al., 2022), sports events (Mousa, 2020; Lowe, 2021) and attending youth camps (Ghosh et al., 2024). Similarly, our conceptual framework and results vindicate how people's degree of favoritism towards specific groups varies based on their closeness (Fong and Luttmer, 2009), salience of shared experiences (McLeish and Oxoby, 2011), or (perceived) similarity (Goeree et al., 2010) to ingroup members. Overall, our results suggest that targeting and reducing interpersonal uncertainty could foster prosocial behavior, bridge animosities, and decrease intergroup conflict.

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Appendix: Proofs

Proof of Theorem 1. For simplicity, whenever possible we will write $E_{v_i \sim f_i}$ simply as E_{f_i} . From the utility expression, we get

$$\begin{aligned}\frac{d}{dx}EU(x) &= E_{f_1, f_2}((v_1 - v_2)U'(v_1x + v_2(100 - x))) \\ \frac{d^2}{dx^2}EU(x) &= E_{f_1, f_2}((v_1 - v_2)^2U''(v_1x + v_2(100 - x))) < 0\end{aligned}$$

as $U'' < 0$ and $f_1, f_2 \geq 0$. $\frac{d^2}{dx^2}EU(x)$ being strictly positive implies that $\frac{d}{dx}EU(x) = 0$ must be obtained at a unique point. Evaluating the first derivative at $x = 50$, we get

$$\frac{d}{dx}EU(x)|_{x=50} = E_{f_1, f_2}(v_1 - v_2)U'(50v_1 + 50v_2) \quad (2)$$

(i) When f_1 and f_2 are identical, then we can also rewrite

$$\begin{aligned}\frac{d}{dx}EU(x) &= E_{v_1 \sim f_2}E_{v_2 \sim f_1}(v_1 - v_2)U'(v_1x + v_2(100 - x)) \\ &= E_{v_2 \sim f_2}E_{v_1 \sim f_1}(v_2 - v_1)U'(v_2x + v_1(100 - x)) \\ &= E_{f_1, f_2}(v_2 - v_1)U'(v_2x + v_1(100 - x))\end{aligned}$$

where the first step integrates v_1 over f_2 and v_2 over f_1 instead, the second step interchanges the names of variables (v_1 and v_2) of integration. The last step interchanges the order of integration. Evaluating the final expression at $x = 50$, we get

$$\frac{d}{dx}EU(x)|_{x=50} = E_{f_1, f_2}(v_2 - v_1)U'(50v_1 + 50v_2) \quad (3)$$

Equations 2 and 3 together imply $\frac{d}{dx}EU(x)|_{x=50} = -\frac{d}{dx}EU(x)|_{x=50} = 0$.

(ii) When f_2 is a mean-preserving spread of f_1 , then there exists a random variable $z \sim f_z$ with zero expectation conditional on any given value of v_1 , such that v_2 has

the same distribution as $v_1 + z$, or in other words, $v_2 \stackrel{d}{=} v_1 + z$. Therefore, we can replace v_2 by a variable $w_1 + z$ where w_1 and v_1 both have identical distribution f_1 .

$$\frac{d}{dx}EU(x) = E_{v_1 \sim f_1} E_{w_1 \sim f_1} E_{z|v_1, w_1} (v_1 - w_1 - z)U'(v_1x + (w_1 + z)(100 - x)) \quad (4)$$

$$\frac{d}{dx}EU(x)|_{x=50} = E_{f_1} E_{f_1} E_{z|v_1, w_1} (v_1 - w_1 - z)U'(50v_1 + 50w_1 + 50z) \quad (5)$$

Because $E_{f_1} E_{f_1}$ is integrating with respect to two identical independent distributions, we can interchange their variable names (w_1 and v_1) in Equation 5:

$$\frac{d}{dx}EU(x)|_{x=50} = E_{f_1, f_1} E_{z|v_1, w_1} (w_1 - v_1 - z)U'(50v_1 + 50w_1 + 50z) \quad (6)$$

Adding equations 5 and 6, and then using law of iterated expectations:

$$\begin{aligned} 2\frac{d}{dx}EU(x)|_{x=50} &= -E_{f_1, f_1} (E_{z|v_1, w_1} 2zU'(50v_1 + 50w_1 + 50z)) \\ &> -E_{f_1, f_1} (E_{z|v_1, w_1} 2zU'(50v_1 + 50w_1)) \\ &= -E_{f_1, f_1} U'(50v_1 + 50w_1) (E_{z|v_1, w_1} 2zf_z(z|v_1)dz) \\ &= 0 \end{aligned}$$

The inequality uses the fact: $zU'(50v_1 + 50w_1 + 50z) < zU'(50v_1 + 50w_1)$ for both $z > 0$ and $z < 0$. The last step follows from the fact that $E_{z|v_1, w_1} z = 0$. Therefore, $\frac{d}{dx}EU(x)|_{x=50} > 0$, and thus, the optimal allocation $x^* > 50$. Next,

$$\begin{aligned} \frac{d}{dx}EU(x)|_{x=100} &= E_{f_1} E_{f_1} E_{z|v_1, w_1} (v_1 - w_1 - z)U'(100v_1) \\ &= E_{f_1} E_{f_1} U'(100v_1) E_{z|v_1, w_1} (v_1 - w_1 - z) \\ &= E_{f_1} E_{f_1} U'(100v_1) (v_1 - w_1) \\ &= E_{f_1} E_{f_1} U'(100w_1) (w_1 - v_1) \\ &= \frac{1}{2} E_{f_1} E_{f_1} [U'(100w_1) (w_1 - v_1) + U'(100v_1) (v_1 - w_1)] \\ &= \frac{1}{2} E_{f_1} E_{f_1} (U'(100w_1) - U'(100v_1))(w_1 - v_1) \\ &< 0 \end{aligned}$$

Step 1 replaces $x = 100$ into the expression of $\frac{d}{dx}EU(x)$ derived at the beginning of the proof. Step 2 uses that $U'(100v_1)$ is independent of z . Step 3 uses $E_{z|v_1, w_1} z = 0$. Step 4 uses the property that v_1, w_1 are drawn i.i.d from f_1 , and hence those two variable names can be interchanged. Step 5 uses the average of the two expressions

from the previous lines. The last step uses the property that U' is decreasing.

As $\frac{d}{dx}EU(x)|_{x=100} < 0$, the concavity of the expression implies that $\frac{d}{dx}EU(x) = 0$ must be obtained at some $50 < x < 100$.

(iii) The first derivative of the objective function, evaluated at x^* , should be zero.

$$E_{f_1, f_2}(v_1 - v_2 - c)U'(x^*v_1 + (100 - x^*)(v_2 + c)) = 0 \quad (7)$$

First, taking the implicit derivative of the last equation w.r.t c , and then re-arranging:

$$E_{f_1, f_2}[-U' + (v_1 - v_2 - c)^2 \frac{dx^*}{dc} U'' + (v_1 - v_2 - c)(100 - x^*)U''] = 0$$

Next, we re-arrange and then bound $\frac{dx^*}{dc}$ in 6 steps as explained below. Under CARA,

$$\begin{aligned} \frac{dx^*}{dc} &= \frac{E_{f_1, f_2} - U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} + \frac{E_{f_1, f_2}(v_1 - v_2 - c)(100 - x^*)U''}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2} - U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} + \frac{E_{f_1, f_2}(v_1 - v_2 - c)(100 - x^*) \times \frac{U''_{100}}{U'_{100}} U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2} - U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} + \frac{(U''_{100})(100 - x^*)}{U'_{100}} \frac{E_{f_1, f_2}(v_1 - v_2 - c)U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2} - U'}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} + \frac{(U''_{100})(100 - x^*)}{U'_{100}} \times \frac{0}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} < 0 \end{aligned}$$

The second step utilizes the assumption of constant absolute risk aversion: $\frac{U''}{U'} = \frac{U''_{100}}{U'_{100}}$, and hence, $U'' = \frac{U''_{100}}{U'_{100}}U'$. The third step simply reorganizes the numerator in the second additive term. The fourth step uses equation 7 to set $E_{f_1, f_2}(v_1 - v_2 - c)U'$ to zero. The last step uses $U' > 0, U'' < 0$.

Under CRRA preferences,

$$\begin{aligned} \frac{dx^*}{dc} &= \frac{E_{f_1, f_2}[-U' + (v_1 - v_2 - c)(100 - x^*)U'']}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2}[-U' - (x^*v_1 + (100 - x^*)(v_2 + c))U'' + 100v_1U'']}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2}[-U' + rU' + 100v_1U'']}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \\ &= \frac{E_{f_1, f_2}[-(1 - r)U' + 100v_1U'']}{-E_{f_1, f_2}(v_1 + c - v_2)^2 U''} \end{aligned}$$

The third step utilizes $r < 1$. In the last expression, the numerator is negative as $r < 1, v_1 \geq 0, U'' < 0$ and the denominator is positive, which concludes the proof. \square

ONLINE APPENDIX

A Additional tables

Table A.1: The influence of changes in interpersonal uncertainty on ingroup favoritism

	<i>Dependent variable:</i>	
	Allocation to ingroup	
	(1)	(2)
Constant (Ingroup preference)	0.741*** (0.132)	0.666*** (0.104)
Interpersonal uncertainty	0.742*** (0.115)	
Higher uncertainty ingroup		-0.856*** (0.144)
Higher uncertainty outgroup		0.629*** (0.135)
Subjects	120	120
Observations	720	720
R ²	0.096	0.095

Notes: The table shows OLS estimates. The dependent variable is the amount allocated to the ingroup (out of \$10), subtracted by five. “Interpersonal uncertainty” is equal to 1, 0, or -1 when the decision has *High uncertainty on ingroup members*, *Group information without uncertainty*, and *High uncertainty on outgroup members* respectively. “Higher uncertainty ingroup” is equal to 1 when the decision has *High uncertainty on ingroup members* (*High uncertainty on outgroup members*) and zero otherwise. Standard errors (in parentheses) are clustered at the subject level. Significance levels: * $p < 0.1$, ** $p < 0.05$ and *** $p < 0.01$.

Table A.2: Structural estimation results (Section 3.5.2)

	Main model (1)
γ	0.374*** (0.016)
b	0.741*** (0.121)
σ	1.881*** (0.046)
a_{IU}	0.739*** (0.152)
a_{ING}	0.336** (0.133)
LL	-1718
Akaike's IC	3445
Bayesian IC	3469

Notes: γ , b , and σ are the CRRA parameter, measure of ingroup preferences, and the standard deviation of the noise term ε respectively. a_{IU} and a_{ING} quantify the importance of interpersonal uncertainty (IU) and ingroup preferences on the optimal allocation choice when both factors are present. See Section 3.5.2 for details. "LL" denotes the maximized Log-Likelihood, "Akaike's IC" is the Akaike's information criterion and "Bayesian IC" the Bayesian information criterion. Significance levels: * $p < 0.1$, ** $p < 0.05$ and *** $p < 0.01$.

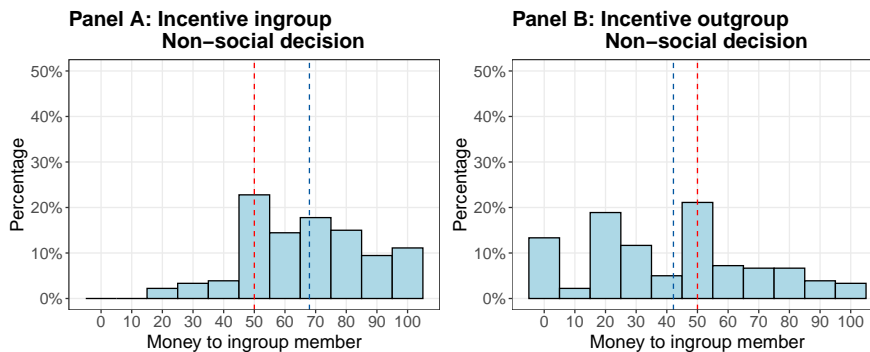
Table A.3: Dictator game allocations under giving setting versus taking setting

	<i>Dependent variable:</i>	
	Allocation to self	
	<i>Social decision</i> (1)	<i>Non-social decision</i> (2)
Constant (Giving setting)	68.050*** (2.719)	66.193*** (2.634)
Taking setting	-27.481*** (3.529)	-12.077*** (3.114)
Order: <i>social</i> decision first	-0.001 (3.545)	2.046 (3.114)
Observations	243	243
R ²	0.201	0.061

Notes: The table shows OLS estimates. The dependent variable is the amount subjects allocate to themselves (out of \$100) in the *Self social* (1) and *Self non-social* decision (2). "Taking setting" is an indicator equal to one if the decision means taking earned money away. "Order: *social* decision first" is an indicator equal to one if subjects faced the *social* before the *non-social* decision. Robust standard errors in parentheses. Significance levels: *p<0.1, **p<0.05 and ***p<0.01.

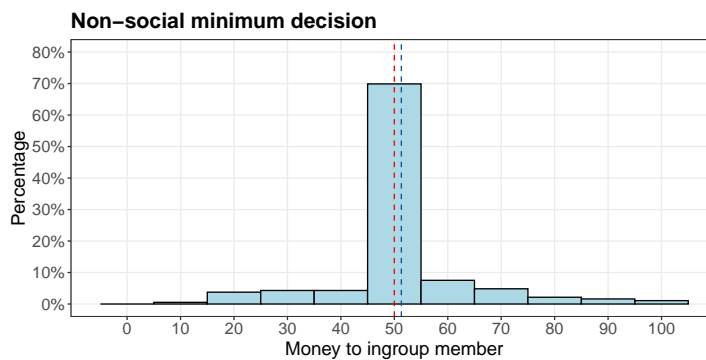
B Additional figures

Figure B.1: Incentive ingroup and Incentive outgroup robustness treatment results



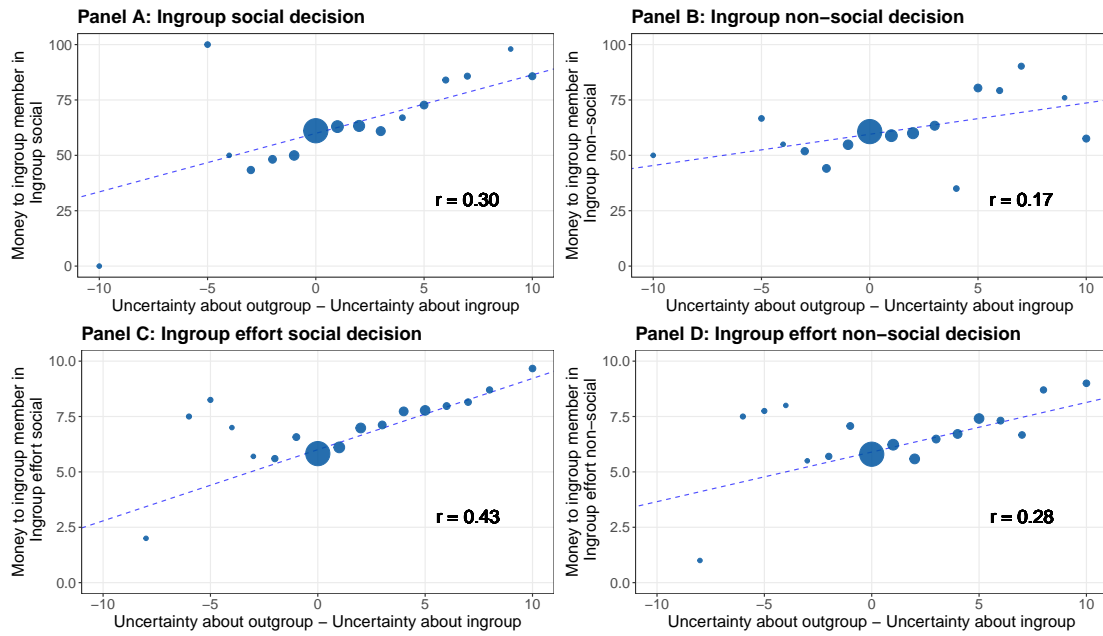
Notes: **Panel A and B:** Histogram of *Incentive ingroup* (Panel A) and *Incentive outgroup* (Panel B) decisions. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to the ingroup member instead of the outgroup member. Subjects incentive is to maximize the weighted sum of the in- and outgroup members WTP. In Panel A, the ingroup receives three times the weight, in Panel B, the outgroup receives three times the weight. The red dotted line denotes the even split benchmark, the blue dotted line the average allocation. For both panels, the binwidth is 10. Decisions are pooled across the three groups (shared hobbies/interests, political views, and religious beliefs), each panel thus displays $n = 180$ decisions by 60 subjects.

Figure B.2: *Non-social minimum* robustness treatment results



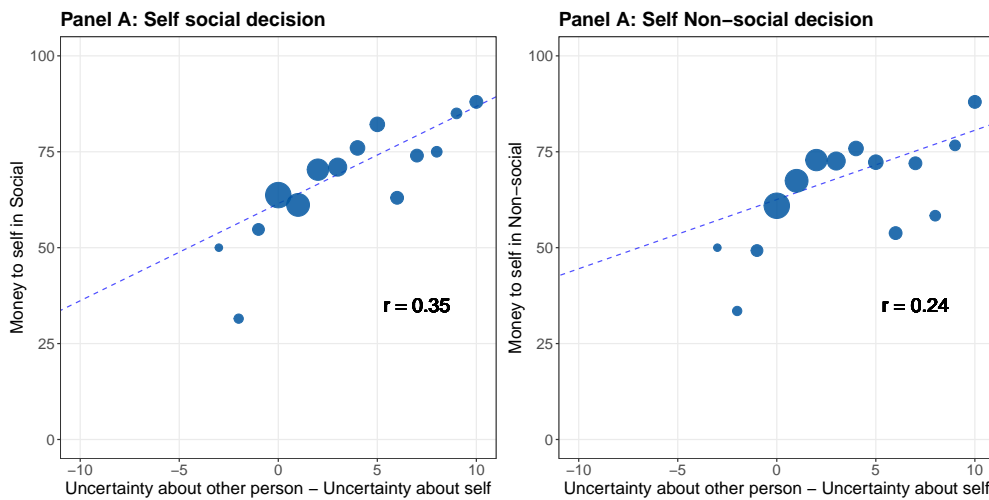
Notes: Histogram of *Non-social minimum* decisions. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to the ingroup member instead of the outgroup member. Subjects incentive is to maximize the minimum of the in- and outgroup member's WTP weighted with subjects' allocation. The red dotted line denotes the even split benchmark, the blue dotted line the average allocation. For both panels, the binwidth is 10. Decisions are pooled across the three groups (shared hobbies/interests, political views, and religious beliefs), each panel thus displays $n = 186$ decisions by 62 subjects.

Figure B.3: Association between interpersonal uncertainty and ingroup versus outgroup decisions



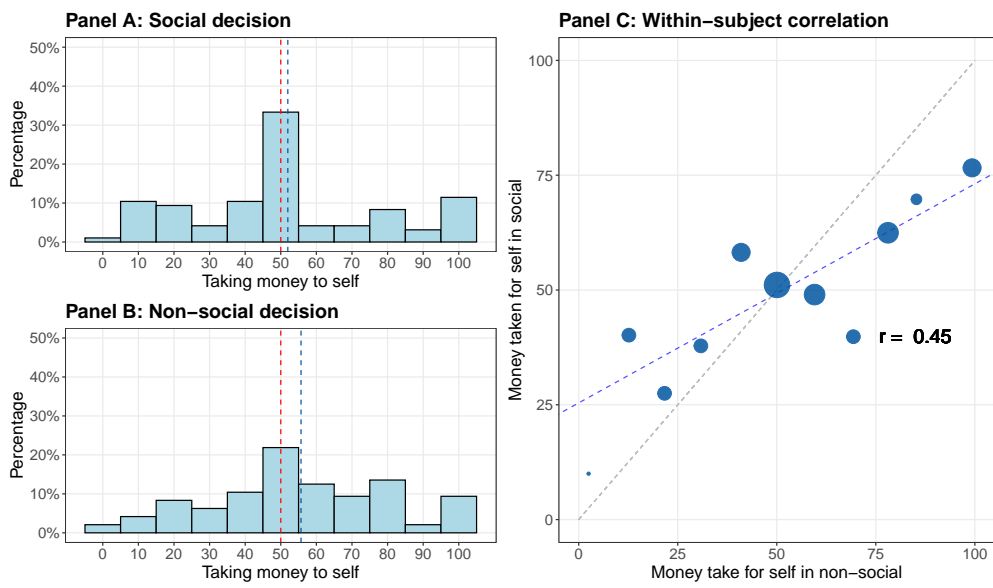
Notes: Binned scatter plot of the association of interpersonal uncertainty with *Ingroup social* decisions in Panel A, with *Ingroup non-social* decisions in Panel B, with *Ingroup effort social* decisions in Panel C, with *Ingroup effort non-social* decisions in Panel D. The x-axis denotes the difference in uncertainty ratings that subjects report about the ingroup and outgroup members gift card value. Higher values indicate more uncertainty about the outgroup's value relative to the ingroup. The y-axis denotes the amount of money that subjects allocate to the ingroup member instead of the outgroup member. In *Ingroup social* and *Ingroup effort social*, the decisions have consequences for the ingroup and outgroup members. In *Ingroup non-social* and *Ingroup effort non-social*, the decisions have consequences only for the subjects, with their payoff depending on the ingroup and outgroup member's (i) WTP for a \$100 gift card in case of *Ingroup non-social* and (ii) WTW for a \$10 bonus payment in case of *Ingroup effort non-social*. The blue dotted line displays the linear fit of a regression of the *social* decisions and *non-social* decisions, respectively, on the difference in interpersonal uncertainty ratings. The correlation coefficient is $r = 0.30$ in *Ingroup social*, $r = 0.17$ in *Ingroup non-social*, $r = 0.43$ in *Ingroup effort social*, and $r = 0.28$ in *Ingroup effort non-social*. The binwidth is 1. Decisions are pooled across the three groups (shared hobbies/interests, political views, and religious beliefs), displaying $n = 357$ decision-pairs by 119 subjects in Panels A and B and $n = 363$ decision-pairs by 121 subjects in Panels C and D.

Figure B.4: Association between interpersonal uncertainty and Dictator game decisions



Notes: Binned scatter plot of the association of interpersonal uncertainty with *Self social* decisions in Panel A and with *Self non-social* decisions in Panel A. The x-axis denotes the difference in uncertainty rating that subjects report about the other person's gift card value and about their own value. The y-axis denotes the amount of gift card money (out of \$100) that subjects allocate to themselves instead of another individual. In Panel A, the decisions have consequences for the subjects and the matched other individuals. In Panel B, the decisions have consequences only for the subjects, with their payoff depending on their and the other individual's WTP for the gift card. The blue line displays the linear fit of a regression of the *Self social* decisions and *Self Non-social* decisions, respectively, on the difference interpersonal uncertainty measure. The correlation coefficient is $r = 0.35$ in *Ingroup social* and $r = 0.24$ in *Ingroup social*. The binwidth is 1. Displayed are $n = 120$ decision-pairs by 120 subjects.

Figure B.5: Giving versus taking results excluding non-takers



Notes: See Figure 4 for details on the variables. Compared to Figure 4, we excluded subjects that take nothing from the other individual in *Taking social*. Thus, displayed are $n = 96$ decisions by 96 subjects.

C Rawlsian preferences under interpersonal uncertainty

In Section 2, we showed that utilitarianism generates patterns of prosocial behavior given certain assumptions on interpersonal uncertainty. This raises the question of whether every commonly used welfare criterion delivers similar patterns under the right parameters given our assumptions. Here, we show that Rawlsian preferences – one of the most discussed welfare criterion – are insensitive to interpersonal uncertainty. Under Rawlsian preferences, only the utility of the least well-off recipient matters. In our context, Rawlsian preferences mean the utility individual i receives from allocating x to the ingroup member and $(100 - x)$ to the outgroup member is $u_{RAWLS} = \min\{v_1x, v_2(100 - x)\}$. As Theorem 2 shows, a decision-maker will then split the money equally independent of interpersonal uncertainty.

Theorem 2. *Suppose individual i has Rawlsian preferences. Then irrespective of i 's risk attitude ($U'' \leq 0$ or $U'' \geq 0$), her optimal allocation is $x^* = 50$, in both the following cases, i) $f_1 = f_2$, and, ii) f_2 is a mean preserving spread of f_1 .*

Proof of Theorem 2. As v_1, v_2 are random variables, i 's expected utility from allocating x to the outgroup is:

$$EU(x) = E_{v_1 \sim f_1, v_2 \sim f_2} \min\{v_1x, v_2(100 - x)\}$$

For any $x \in [0, 50) \cup (50, 100]$,

$$\min\{v_1(100 - x), v_2x\} + \min\{v_1x, v_2(100 - x)\} \leq v_1(100 - x) + v_1x = 100v_1$$

with strict inequality whenever $v_1 \neq v_2$.²⁴

Similarly, $\min\{v_1(100 - x), v_2x\} + \min\{v_1x, v_2(100 - x)\} \leq 100v_2$ with strict inequality whenever $v_1 \neq v_2$. Putting these two inequalities together, we get

$$\min\{v_1(100 - x), v_2x\} + \min\{v_1x, v_2(100 - x)\} \leq \min\{100v_1, 100v_2\}$$

with strict inequality whenever $v_1 \neq v_2$. Next, using $f_1 = f_2$,

$$\begin{aligned} EU(x) &= E_{v_1 \sim f_1, v_2 \sim f_2} \min\{v_1x, v_2(100 - x)\} \\ &= E_{v_1 \sim f_1, v_2 \sim f_2} \min\{v_1(100 - x), v_2x\} \end{aligned}$$

²⁴If $x = 50$, then strict inequality does not hold under $v_1 < v_2$.

Therefore, for any $x \in [0, 50) \cup (50, 100]$,

$$\begin{aligned} EU(x) &= \frac{1}{2} \times E_{v_1 \sim f_1, v_2 \sim f_2} (\min\{v_1 x, v_2(100 - x)\} + \min\{v_1(100 - x), v_2 x\}) \\ &< \frac{1}{2} \times E_{v_1 \sim f_1, v_2 \sim f_2} \min\{100v_1, 100v_2\} \\ &= E_{v_1 \sim f_1, v_2 \sim f_2} \min\{50v_1, 50v_2\} \end{aligned}$$

The first inequality becomes strict as $v_1 \neq v_2$ with positive probability in the integration. This proves part (i), and a similar proof works for part (ii) after v_2 is replaced with $w_1 + z_1$ like in the proof of Theorem 1.

□

D Within-subject analyses

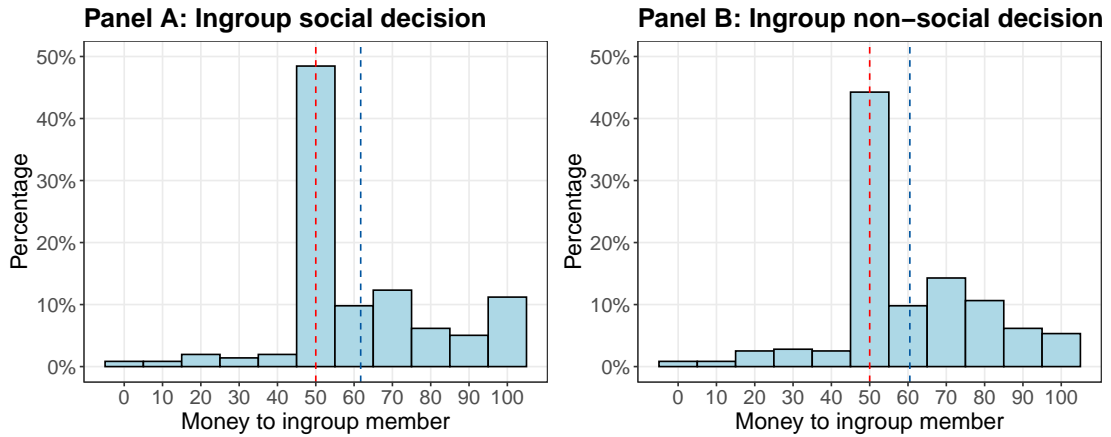
The results covered in the main text were obtained using a between-subject design, where we only used the first set of decisions each subject faced. In the following, we repeat our analyses using all of the subjects' decisions. In general, our between-subject results replicate well in the within-subject analyses.

D.1 Ingroup versus outgroup paradigm main results

Ingroup social decisions. In the within-subject case, subjects allocate on average \$57.48 if their ingroup members share the same interests/hobbies, \$67.81 if political views are shared, and \$59.88 if religious beliefs are shared. In all three cases, we can reject the hypothesis of no ingroup favoritism ($p < 0.001$, one-sample Wilcoxon tests). Figure D.1 panel A displays the distribution pooled over the three decisions. In 52% of the decisions, subjects display ingroup favoritism by allocating strictly more than 50% to the ingroup. Outgroup favoritism is found in 9% of decisions, and in the remaining 39%, subjects allocate 50/50. In total, 76% of subjects display ingroup favoritism in at least one decision.

Ingroup non-social decisions. In the *Ingroup non-social* decisions, subjects allocate on average \$58.47 to their ingroup members sharing the same interests/hobbies, \$64.00 if political views are shared and \$58.97 if religious beliefs are shared. Again,

Figure D.1: Ingroup within-subject



Notes: See Figure 1 for details on the variable definitions. Compared to Figure 1, this Figure displays the results using all decisions from subjects. Each panel thus display $n = 357$ decisions by 119 subjects.

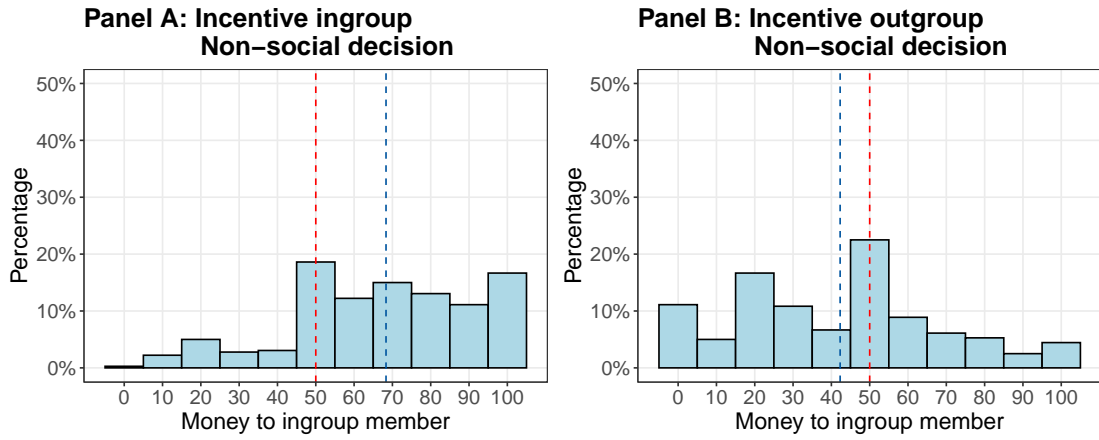
we can reject the hypothesis of no ingroup favoritism ($p < 0.001$, one-sample Wilcoxon tests) in all three cases. Figure D.1 panel B displays the distribution. In 55% of the decisions, subjects display ingroup favoritism by allocating strictly more than 50% to the ingroup. Outgroup favoritism is found in 12% of decisions, and in the remaining 32%, subjects allocate 50/50.

Comparing Ingroup social and non-social. Comparing average ingroup allocations between *Ingroup social* and *non-social* within-subject reveals that we cannot reject equality in all three cases ($p = 0.59$ for hobbies/interests, $p = 0.20$ for political views, $p = 0.94$ for religious beliefs, paired Wilcoxon tests). Further, we cannot reject that the distributions of allocations are equal ($p = 0.30$ for hobbies/interests, $p = 0.23$ for political views, $p = 0.99$ for religious beliefs, Kolmogorov-Smirnov test). The same holds when decisions are pooled across the three domains for additional statistical power ($p = 0.40$, Kolmogorov-Smirnov test).

D.2 Ingroup versus outgroup setting incentive robustness

Table D.1 displays the treatment of *Outgroup incentive* relative to *Ingroup incentive* effects separately for the within-subject and between-subject effects pooled across the

Figure D.2: Ingroup incentive within-subject



Notes: See Figure B.1 for details on the variable definitions. Compared to Figure B.1, this Figure displays the results using all decisions from subjects. Each panel thus displays $n = 360$ decisions by 120 subjects.

three groups. As displayed, the effect is similar in both the within- and between-subject comparison. Regarding the within-subject effects in the social groups individually, when the ingroup is incentivized, average ingroup allocations increase from \$58.47 to \$67.22 for hobbies/interests ($p < 0.001$, unpaired Wilcoxon tests), from \$64.00 to \$72.57 for political views ($p < 0.001$), and from \$58.97 to \$65.22 for religious beliefs ($p = 0.01$) compared to *Ingroup non-social*. Conversely, in *outgroup incentive*, allocations to the ingroup decrease to \$37.76 for hobbies/interests, to \$46.89 for politics and to \$42.21 for religious beliefs (all three $p < 0.001$). As in the between-subject comparison, we again see outgroup favoritism in the *Outgroup incentive* decisions. The pooled average is \$42.29, which is significantly different from the even split ($p < 0.001$, one-sample Wilcoxon test). See Figure D.2 for the distributions, which once again show that the shift in average giving is driven by shifts in the distributions.

D.3 Self versus others setting main results

Self social decision. In the within-subject case of the *Self social* decision, subjects allocate on average \$68.05 to themselves, thus displaying significant self-regarding

Table D.1: Treatment effect of the incentive treatment in the ingroup versus outgroup setting

	<i>Dependent variable:</i>	
	Allocation to ingroup member	
	Within-subject (1)	Between-subject (2)
<i>Outgroup incentive</i>	−26.047*** (3.163)	−25.817*** (3.547)
Constant (<i>Ingroup incentive</i>)	68.333*** (1.937)	67.978*** (2.083)
Subjects	120	120
Observations	720	360
R ²	0.211	0.231

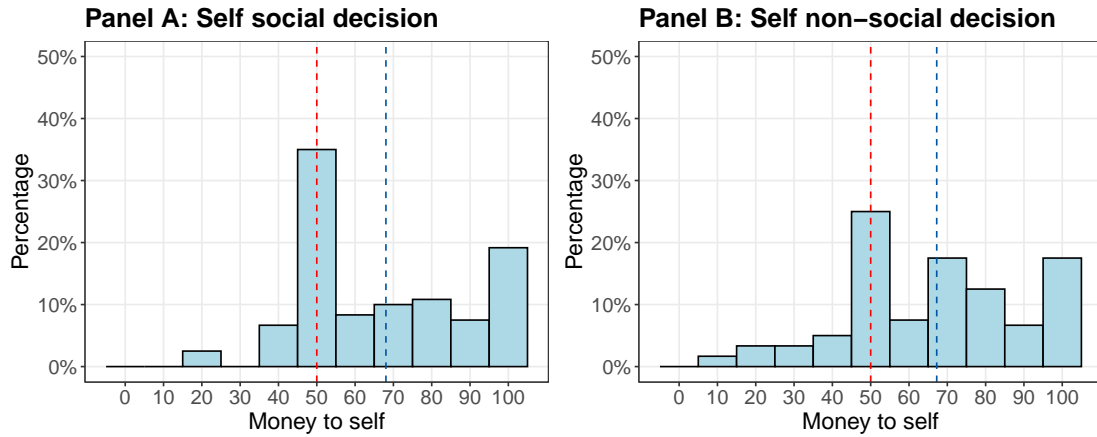
Notes: The table shows OLS estimates. The dependent variable in columns (1) and (2) is the amount subjects allocate to themselves (out of \$100) in the *Ingroup incentive* and *Ingroup incentive* treatments. “Outgroup incentive” is a dummy equal to one if the incentive for the decision gave three times the weight on the outgroup member’s WTP, and equal to zero if the incentive gave three times the weight on the ingroup member’s WTP. In column (1), all decisions are used, in (2) only the first decisions. Standard errors (in parentheses) are clustered at the subject level. Significance levels: * $p < 0.1$, ** $p < 0.05$ and *** $p < 0.01$.

behavior relative to the equal split ($p < 0.001$, one-sample Wilcoxon test). Figure D.3 panel A displays the distribution. In total, 62% of subjects allocate more money to themselves, 9% allocate more to the other person, and 29% implement the 50/50 split.

Self non-social decision. In the *Self non-social* decision, subjects allocate on average \$67.23 to themselves, again displaying significant self-regarding behavior ($p < 0.001$, one-sample Wilcoxon tests). Figure D.3 panel B shows the distribution. In total, 66% of subjects allocate more money to themselves, 13% allocate more to the other person, and 21% implement the 50/50 split.

Comparing Self social and non-social. In the between-subject comparison, we also cannot reject equality of average allocations between *Self social* and *non-social* ($p = 0.69$, paired Wilcoxon tests). Similarly, we cannot reject that the pooled distributions

Figure D.3: Dictator game between subject



Notes: Histogram of the *Self social* (Panel A) and *Self non-social* (Panel B) decision. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to themselves instead of another individual. The red dotted line denotes the even split benchmark, the blue dotted line the average allocation. In *Self social* (Panel A), the decision has consequences for the subjects and the other individual. In *Self non-social* (Panel B), the decision has consequences only for the subjects, with their payoff depending on their and the other individual's WTP for the gift card. For both panels, the bin-width is 10. Both panels display $n = 240$ decisions by 120 subjects.

are equal ($p = 0.95$, Kolmogorov-Smirnov test).

E Analyzing order effects

A potential concern for the validity of the within-subject results for the *social* and *non-social* decisions (Appendix D) is contagion across conditions. As subjects facing the first set of decisions were not aware that a second set would follow, this naturally cannot influence our between-subject analysis presented in the main paper that only uses the first set of decisions. However, subjects may adjust their choice in the subsequent *non-social* decisions to mimic the *social* decisions, potentially biasing the individual-level analyses. Such adjustment could lead to artificially high similarity between the two decisions, and thus artificially high correlations. Because we randomized the order of decisions, we can directly assess this concern by testing for order effects. Overall, we find no evidence that the order influences subjects' behavior, as we show in the following in detail.

E.1 Ingroup versus outgroup paradigm

For the *Ingroup non-social* decisions, the pooled average allocations to the ingroup are \$60.89 when elicited before, and \$60.10 when elicited after the *social* decisions. For hobbies/interests *Ingroup non-social* the averages are \$56.86 and \$59.95 ($p = 0.58$, unpaired Wilcoxon test), for political views \$65.02 and \$63.06 ($p = 0.61$), and for religious beliefs \$60.81 and \$57.29 ($p = 0.15$). Thus, the averages are invariant to the order. We also cannot reject the null that distributions are invariant to the order ($p = 0.61, p = 0.31, p = 0.22$, Kolmogorov-Smirnov tests). Moving to the *Ingroup social* decisions, average allocations are \$63.38 when *Ingroup social* is elicited first, and \$59.92 when elicited after the *non-social* decisions. Again, averages and distributions generally do not differ significantly. For hobbies/interests *Ingroup social* the averages are \$57.48 and \$57.47 ($p = 0.54$, unpaired Wilcoxon test), for political views \$71.05 and \$64.28 ($p = 0.26$), and for religious beliefs \$61.61 and \$58.00 ($p = 0.99$). We also cannot reject the null that distributions are invariant to the order for hobbies/interests and religious beliefs ($p = 0.27$ and $p = 0.32$, Kolmogorov-Smirnov tests), with the only exception being political views ($p = 0.01$).

E.2 Self versus other paradigm

In the case of the *Self non-social* decisions, subjects allocate \$64.12 to themselves when the decision is before the *Self social* decision, and \$70.25 when the decision comes afterward, an insignificant difference ($p = 0.12$, unpaired Wilcoxon test). In the case of the *Self social* decisions, subjects allocate \$69.05 to themselves when the decision is before the *Self non-social* decision, and \$67.02 when the decision comes afterward, again an insignificant difference ($p = 0.61$). In addition, we can reject the null that the distributions are invariant to the order at the 5% level for *Self non-social* ($p = 0.08$ Kolmogorov-Smirnov tests) and at any conventional level for *Self social* ($p = 0.54$ Kolmogorov-Smirnov tests).

E.3 Giving versus taking paradigm

In the case of the *Taking non-social* decisions, subjects allocate \$56.00 to themselves when the decision is before the *Taking social* decision, and \$54.05 when the decision comes afterward, an insignificant difference ($p = 0.49$, unpaired Wilcoxon test). In the case of the *Taking social* decisions, subjects allocate \$39.52 to themselves when the decision is before the *Taking non-social* decision, and \$41.51 when the decision comes afterward, again an insignificant difference ($p = 0.52$). In addition, we cannot reject the null that the distributions are invariant to the order both for *Taking non-social* ($p = 0.45$ Kolmogorov-Smirnov tests) and *Taking social* ($p = 0.73$ Kolmogorov-Smirnov tests).

F Further details on the robustness treatments of Section (Section 3.3)

This section provides more details on the results of the robustness treatments that are described in Section 3.3.

F.1 Further results from Section 3.3.1

We start with the between-subject comparison of *Outgroup incentive* versus *Ingroup incentive* by focusing on behavior in the first-assigned choices. We find that subjects respond to changes in the induced utilitarian incentives: Compared to an average ingroup giving of \$60.89 across our social groups in *Ingroup non-social* (Section 3.2), subjects give on average \$42.16 to the ingroup member in *Outgroup incentive* and \$67.98 in *Ingroup incentive*, a significant difference (both comparisons $p < 0.001$, unpaired Wilcoxon test). In particular, behavior switches to outgroup favoritism in *Outgroup incentive*, as the average is significantly smaller than the even split ($p < 0.001$, one-sample Wilcoxon test). Regarding the distributions, 54% of subjects in *Outgroup incentive* and 10% in *Ingroup incentive* displaying outgroup favoritism, while 33% and 72% display ingroup favoritism, respectively. This significant shift in the distributions ($p < 0.01$, Kolmogorov-Smirnov test) shows that the changes in the averages are not

driven by a minority of subjects, but a substantial fraction. See Appendix Figure B.1 for the histograms. The within-subject analysis (see Appendix D.2 for details) reveals that subjects change their behavior in 81% of decisions following the incentive change between the two treatments. Accordingly, in 19% of decisions are subjects unresponsive to changes in the incentives, indicating inattention or confusion. Comparing the behavior in these situations to the main experiment's *Ingroup social*, both average ingroup giving (\$60.84 compared to \$63.38) and the fraction of choices displaying ingroup favoritism (34% of decisions compared to 46%) is lower, while 50/50 splits are more frequent (60% to 46%).

Next, we turn to the *Ingroup non-social minimum* treatment. As predicted, we find that favoritism in either direction is eliminated under the Rawlsian incentive. On average, subjects allocate \$51.31 to the individual sharing their interests/hobbies, \$52.85 to the individual sharing political views and \$49.77 to the individual sharing religious beliefs. Hence, the treatment did not only significantly reduce ingroup favoritism relative to *Ingroup non-social* (in all three cases $p < 0.01$, unpaired Wilcoxon tests), but eliminated it altogether, as we can no longer reject that average ingroup giving is different from the 50/50 split ($p = 0.31$, $p = 0.13$ and $p = 0.95$ respectively, one-sample Wilcoxon tests). See Appendix Figure B.2 for the distribution, which further demonstrates that subjects respond strongly to the induced incentives in the expected direction: the percentage of decisions that implement exactly a 50/50 split increases from 32% in *ingroup non-social* to 58% *Ingroup non-social minimum*.

These results provide evidence against limited attention or confusion driving our results of the previous section. Changing a single number in the Utilitarian incentive formula reverses the direction of favoritism from ingroup to outgroup favoritism, while moving to Rawlsian incentives eliminates any favoritism. Moreover, in those cases where choices indicate inattention or confusion about the incentives, ingroup favoritism is, if anything, less prevalent relative to the main experiment.

F.2 Further results from Section 3.3.2

This section provides details on the results of the *Ingroup effort* treatment (Section 3.3.2). The treatment contains a *Ingroup effort social* and *Ingroup effort non-social* de-

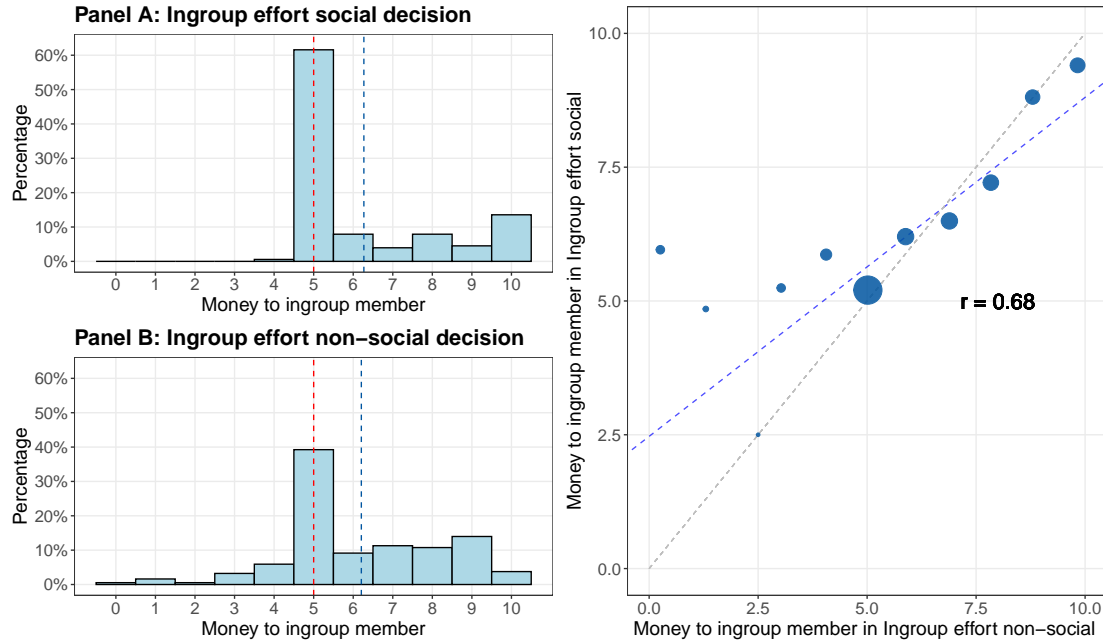
cision. In the *Ingroup effort social* decisions, subjects allocate on average \$6.27 to the ingroup member, i.e., 63% of the endowment, almost identical to the 61% found in *Ingroup social*. Ingroup favoritism is found in 41% of decisions (46% in *Ingroup social*) while subjects display outgroup favoritism in 2% (8%) of decisions. The remaining 57% (46%) of decisions are 50/50 allocations. In *Ingroup effort non-social*, the average giving to the ingroup member is 62% of the endowment compared to 61% in *Ingroup non-social*, with 57% (61%) of decisions showing ingroup-favoritism, 12% (11%) showing outgroup-favoritism, and the remaining 31% (28%) decisions are 50/50 allocations. As before, we also find a strong within-subject correlation of 0.68 between *Ingroup effort social* and *Ingroup effort non-social*. See Figure F.1 for the corresponding between-subject histograms and the within-subject bin-scatter plot. To summarize, both *Ingroup effort social* and *Ingroup effort non-social* closely replicate the *Ingroup social* and *Ingroup non-social* decisions not only in terms of directions of patterns, but also in quantitative magnitude. Accordingly, our results do not depend on the use of gift card money as medium of allocation in the *Social* decisions and as part of the incentive in the *Non-social* decisions.

G Ingroup information decision

In this section, we discuss the design and results of an additional decision that was added to the end of the *Ingroup uncertainty* treatment described in Section 3.4.2.

Design. After making the seven decisions that exogenously informed subjects which group had ingroup affiliation and/or which group had a higher WTW-variance, subjects made an eighth and final allocation decision. In this decision, they were not given any information that distinguished one group from the other ex-ante. They were only informed that the two groups were different: (i) both recipients of one group (which group was unspecified) shared their hobbies/interests (or shared their political views), while both recipients of the other group did not, and (ii) both recipients of one group (which group was unspecified) had a WTW of 12, while the two recipients of the other group had a WTW of 4 and 22. However, ex-ante, they did not know which group had the ingroup members, or which group had a lower

Figure F.1: Results effort ingroup versus outgroup decisions



Notes: **Panel A and B:** Histogram of *Ingroup effort social* (Panel A) and *Ingroup effort non-social* (Panel B) decisions. The x-axis denotes the amount of money (out of \$10) that subjects allocate to the ingroup member instead of the outgroup member. The red line denotes the even split benchmark, the blue line the average allocation. In Panel A, the decisions have consequences for the ingroup and outgroup members. In Panel B, the decisions have consequences only for the decision-makers, with their payoff depending on the ingroup and outgroup member's willingness-to-work on real effort tasks for the money. **Panel C:** Binned scatter plot of *Ingroup effort social* and *Ingroup effort non-social* decisions. The blue line displays the linear fit of regressing *Ingroup effort social* on *Ingroup effort non-social* decisions. The correlation coefficient is $r = 0.53$. For all three panels, the binwidth is 1. Decisions are pooled across the three groups (shared hobbies/interests, political views, and religious beliefs), displaying $n = 177$ decisions by 59 subjects in Panel A, $n = 186$ decisions by 62 subjects in Panel B, and $n = 363$ decision-pairs by 121 subjects in Panel C.

variance in WTW, or if ingroup members had lower/ higher WTW variance.

Subjects then could choose to learn one of the two dimensions along which the groups differed. That is, they could either learn which group contained only ingroup members and which contained only outgroup members, or, they could learn which group had WTW variation and which had not. We balanced the presentation of both pieces of information by randomizing the order in which the information was introduced and displayed for the choice. Which information subjects choose reveals which factors are of primary importance in their decision process.

Results. We find that 81% of subjects choose to learn about the WTW information, and only 19% choose to learn the ingroup-outgroup information. Accordingly, in our setting, subjects prefer receiving information about interpersonal uncertainty over information about social group membership, indicating its relevance for decisions.

Information choice type analysis. We further use subjects' preference towards receiving information about the recipients' WTWs or group affiliations to validate our type categorization that we develop in Section 3.5. We compare the fraction of subjects choosing the WTW information instead of the group affiliation information across our four behavioral types. In total, 92% of subjects who respond to uncertainty but display no group preference choose the WTW information. This fraction decreases to 78% for those who respond to uncertainty and reveal a group preference, and decreases further to 50% for those who reveal a group preference but do not respond to uncertainty. Hence, our categorization predicts subjects' information choices in the expected direction.

H Self versus others setting incentive treatment

Design. As in the ingroup case, we vary the incentive subjects face when making the *Self non-social* decisions. In *Self incentive*, the weight on the DM's own WTP is three times as high as the other individuals WTP. The DM's payoff thus becomes:

$$\Pi(x_{self}, x_{other}) = 3 \cdot x_{self} \cdot WTP_{self}/100 + x_{other} \cdot WTP_{other}/100$$

In *Other incentive* we increase the weight put on the other individual's WTP to be three times as high as the DM's WTP:

$$\Pi(x_{self}, x_{other}) = x_{self} \cdot WTP_{self}/100 + 3 \cdot x_{other} \cdot WTP_{other}/100$$

Results. Inducing these incentives changes people's behavior in the *non-social* decision. See Table H.1 for the within-subject and between-subject treatment effect. In both cases lead the change in incentives to a significant change in the amount subjects allocate to themselves, they allocate \$19.50 in the within and \$22.08 in the between-subject comparison less to the themselves when the incentives are higher for the other participant. Figure H.1 displays the distributions in the between-subject

case. The fraction of subjects allocating more than 50% of the endowment to themselves increases from 33% in *Other incentive* to 63% in *Self incentive*, while the fraction of subjects allocating more money to the other participant decreases from 50% to 22%. The distributions are significantly different from each other ($p < 0.001$, Kolmogorov-Smirnov test).

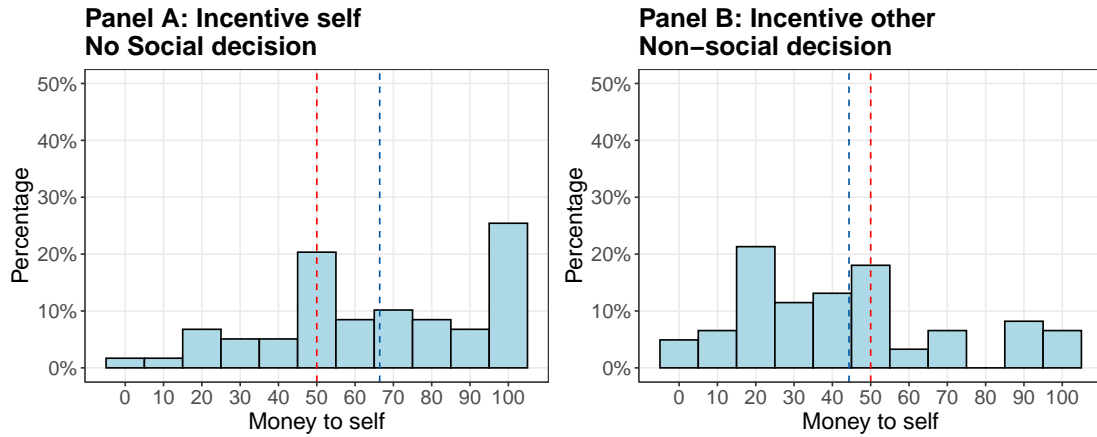
Using the within-subject comparison shows that 83% of subjects change their allocation behavior between *Self incentive* and *Other incentive*. Among those 17% subjects that are unresponsive to the incentive change, 33% allocate more to themselves, a substantially lower fraction than the 58% in the main *Self non-social* case. In total, 38% choose the equal split, and 29% allocate more to the other participant. Taking the behavior of these unresponsive subjects as indicative of inattention or confusion, it appears that such factors are associated with subjects allocating less to themselves. This result thus provides suggestive evidence that our replication of significantly more self-giving in *Self social* using the *Self non-social* decisions is not driven by inattentive or confused subjects.

I Validation of the uncertainty measure

Our self-reported interpersonal uncertainty measure is intended to proxy whether subjects perceive higher interpersonal uncertainty of one group over another, as defined in Definition 1 in Section 2. However, it could be the case that subjects instead only report their perception about mean differences or concepts unrelated to uncertainty. In this section, we validate that our measure is indeed sensitive to those changes in interpersonal uncertainty captured by our definition. To do so, we provide subjects with two objective WTP distributions, one being a mean-preserving spread of the other, and investigate the impact on the answers subjects give to our measure.

Design. At the end of the *Ingroup belief measurement* and *Self/other belief measurement* treatments, we showed subjects two figures. Each figure displayed a frequency distribution of the WTP values of 100 fictitious individuals. In one, 50 individuals had a WTP of \$86 and 50 a WTP of \$88 (low variance distribution). In the other were 10 individuals for each of the 10 values between \$78 and \$96 (high variance dis-

Figure H.1: Self versus other incentive



Notes: Histogram of *Self incentive* (Panel A) and *Other incentive* (Panel B) decisions. The x-axis denotes the amount of gift card money (out of \$100) that subjects allocate to themselves instead of another individual. Subjects incentive is to maximize the weighted sum of their own and another individual's WTP. In Panel A, subjects own WTP receives three times the weight, in Panel B, the other individual's WTP receives three times the weight. The red dotted line denotes the even split benchmark, the blue dotted line denotes the average allocation. For both panels, the binwidth is 10. Only the first decision is used for each subject. Panel A displays $n = 59$ decisions by 59 subjects, Panel B displays $n = 61$ decisions by 61 subjects.

tribution). See Figure I.1 for the figures shown to subjects. We also provided these values to subjects in text format below the figures. The high variance distribution is a mean-preserving spread of the low variance one, having the same mean but a lower variance. For each figure, subjects were asked the following about the group displayed in the figure: “Suppose we randomly pick one of the 100 people from this group. How certain are you about how much the randomly chosen person would value the Amazon gift card money?” Subjects could respond on an 11-point Likert scale from *Very uncertain* to *Very certain*, and we re-code the variable so that higher values indicate higher uncertainty. The text and measurement thus closely mirror our self-reported interpersonal uncertainty measure.

Results. We find that subjects report different uncertainty across the two distributions. On average, they report an uncertainty of 3.46 Likert-scale points for the low variance distribution, and an uncertainty of 5.80 points for the high variance distribution, a significant difference ($p < 0.001$, paired Wilcoxon-test). On the individual

Table H.1: Treatment effect of the incentive treatment in the self versus other setting

	<i>Dependent variable:</i>	
	Allocation to self	
	Within-subject (1)	Between-subject (2)
<i>Other incentive</i>	−19.500*** (3.468)	−22.079*** (5.045)
Constant (<i>Self incentive</i>)	65.625*** (2.563)	66.424*** (3.592)
Subjects	120	120
Observations	240	120
R ²	0.108	0.140

Notes: The table shows OLS estimates. The dependent variable is the amount subjects allocate to themselves (out of \$100) in the *Other incentive* and *Self incentive* treatments. “Other incentive” is an indicator equal to one if the incentive for the decision gave three times the weight on the other person’s WTP, and zero if the incentive gave three times the weight on the subject’s own WTP. In column (1), all decisions are used, in (2) only the first decisions. Standard errors (in parentheses) are clustered at the subject level. Significance levels: * $p < 0.1$, ** $p < 0.05$ and *** $p < 0.01$.

level, 74% of subjects report a higher uncertainty for the high variance distribution compared to the low variance distribution, 13% report no difference, and the remaining 13% report more uncertainty for the low variance distribution. Thus, subjects are sensitive to changes in WTP distributions in the expected direction.

J Research transparency

All experiments covered in the paper were preregistered at aspredicted.org. The pre-registrations include details on the experimental design, the planned sample size, exclusion criteria, hypotheses, and the main analyses. Table J.1 provides an overview over the treatments and links to the respective pre-registrations.

Our experimental implementation followed closely the pre-registration. In partic-

Figure I.1: Distributions shown to subjects

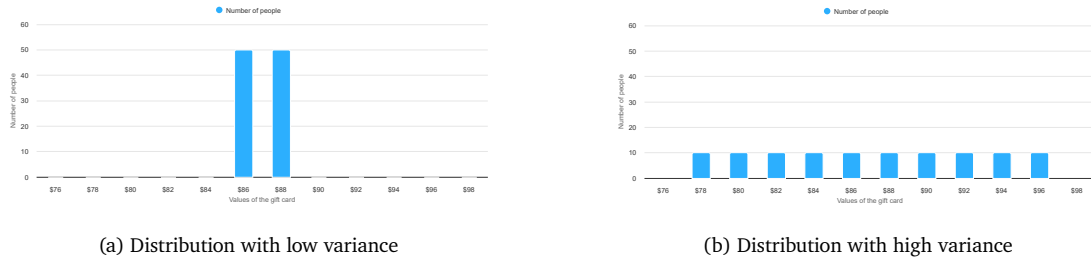


Table J.1: Overview over treatments

Label	N	Covered in	Link to preregistration
<i>Ingroup social & Ingroup non-social</i>	119	Section 3	https://aspredicted.org/H81_KQ5
<i>Ingroup incentive & Outgroup incentive</i>	120	Section 3.3	https://aspredicted.org/H81_KQ5
<i>Ingroup minimum</i>	62	Section 3.3	https://aspredicted.org/J7H_W8R
<i>Ingroup effort social & Ingroup effort non-social</i>	121	Section 3.3	https://aspredicted.org/53G_PNJ
<i>Ingroup uncertainty</i>	120	Section 3.4	https://aspredicted.org/53G_PNJ
<i>Self social & Self non-social</i>	120	Section 4	https://aspredicted.org/ZMF_CD9
<i>Self incentive & Other incentive</i>	120	App. Section H	https://aspredicted.org/ZMF_CD9
<i>Self taking social & Self taking non-social</i>	123	Section 5	https://aspredicted.org/RT4_TQB
<i>Ingroup belief measurement & Self/other belief measurement</i>	120	Sections 3.4, 4.3, 5.2	https://aspredicted.org/T7X_747

ular, we implemented the experimental design and sample size exactly as specified in the pre-registration. Similarly, we employed the exclusion criteria as pre-registered: we specified to exclude any subject who did not complete the experiment. This led to the exclusion of 22 subjects in the *Ingroup social* and *Ingroup non-social* treatments, 28 in the *Ingroup incentive* and *Outgroup incentive*, 6 in the *Ingroup minimum*, 13 in the *Ingroup effort social* and *Ingroup effort non-social*, 15 in *Ingroup uncertainty*, 22 in *Self social* and *Self non-social*, 23 in *Self incentive* and *Other incentive*, 25 in *Self taking social* and *Self taking non-social*, 3 in *Ingroup belief measurement* and 1 in *Self/other belief measurement*. The sample sizes reported in Table J.1 are the final sample sizes used in all analyses of the paper after excluding the previously mentioned numbers of subjects.

J.1 Deviations from the pre-registration

The pre-registrations for the *Ingroup social* and *Ingroup non-social* treatments as well as the *Self social* and *Self non-social* treatments contain another set of treatments labeled *Group info* and *Self info*. These treatments are not part of this paper and their results are available upon request because the design is superseded by the *Ingroup uncertainty* experiment.²⁵ The analyses contained in Section 3.5 were pre-registered as exploratory analyses without specifying any details.

K Experimental instructions

The instructions of all experiments can be found in the following Open Science Framework (OSF) repository:

https://osf.io/tcp3d/?view_only=27899531990048e4b608d64b13528236

²⁵The omitted treatments show that providing subjects with information on the WTP of the recipients significantly changes their allocation behavior both in *Ingroup social* and *Self social*. However, in contrast to the *Ingroup uncertainty* experiment, this information manipulation does not directly manipulate interpersonal uncertainty and is potentially confounded by experimenter demand effects.